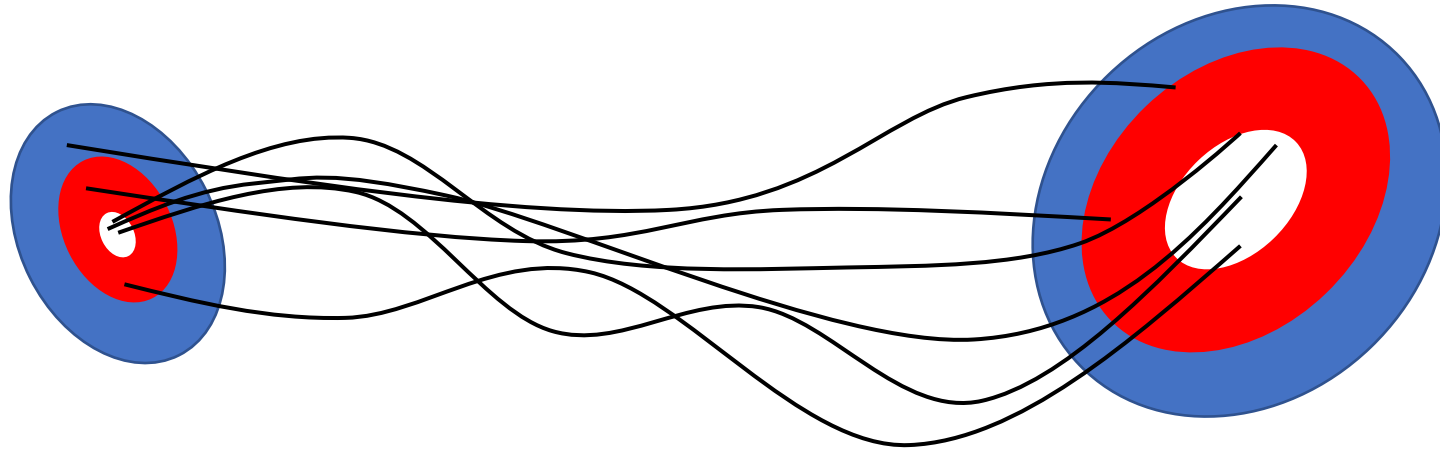


The Ensemble Kalman filter



Part II: Practicalities

Yumeng Chen (Based on notes by Ross Bannister, Alison Fowler, and Ivo Pasmans)

Quick recap

- Recall that the ensemble covariance matrix at time step k is given by with N ensemble members.

$$\mathbf{P}_k^f = \frac{1}{N-1} \mathbf{X}_k'^f (\mathbf{X}_k'^f)^T = \frac{1}{N-1} \sum_{i=1}^N \left(\mathbf{x}_k^{(i),f} - \bar{\mathbf{x}}_k^f \right) \left(\mathbf{x}_k^{(i),f} - \bar{\mathbf{x}}_k^f \right)^T$$

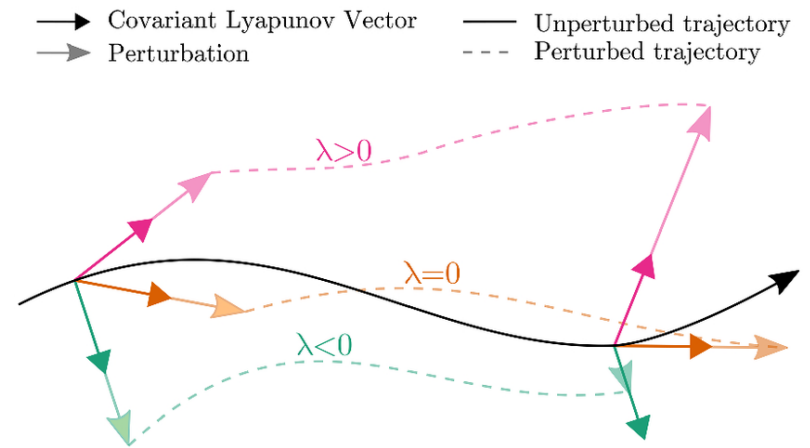
- We can write a similar expression for the analysis error covariance matrix in terms of the analysis perturbations

$$\mathbf{P}_k^a = \frac{1}{N-1} \mathbf{X}_k'^a (\mathbf{X}_k'^a)^T = \frac{1}{N-1} \sum_{i=1}^N \left(\mathbf{x}_k^{(i),a} - \bar{\mathbf{x}}_k^a \right) \left(\mathbf{x}_k^{(i),a} - \bar{\mathbf{x}}_k^a \right)^T$$

where $\mathbf{X}' \in \mathbb{R}^{n \times N}$ is the **perturbation matrix**.

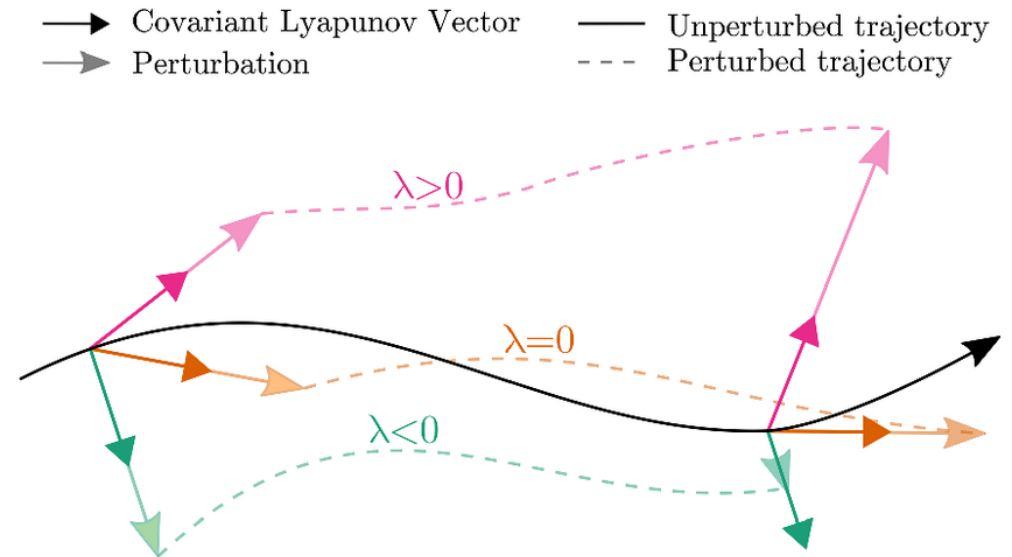
How much ensemble is sufficient?

- A dynamical system with n number of states can have up to n different modes. Typically, only $N < n$ modes can lead to increased error from small errors in the initial condition.
- In the weather systems, unstable convective storms and stable high pressure systems (imagining “anticyclonic gloom”) can occur simultaneously.
- Mathematically, these modes are quantified by Lyapunov exponents.



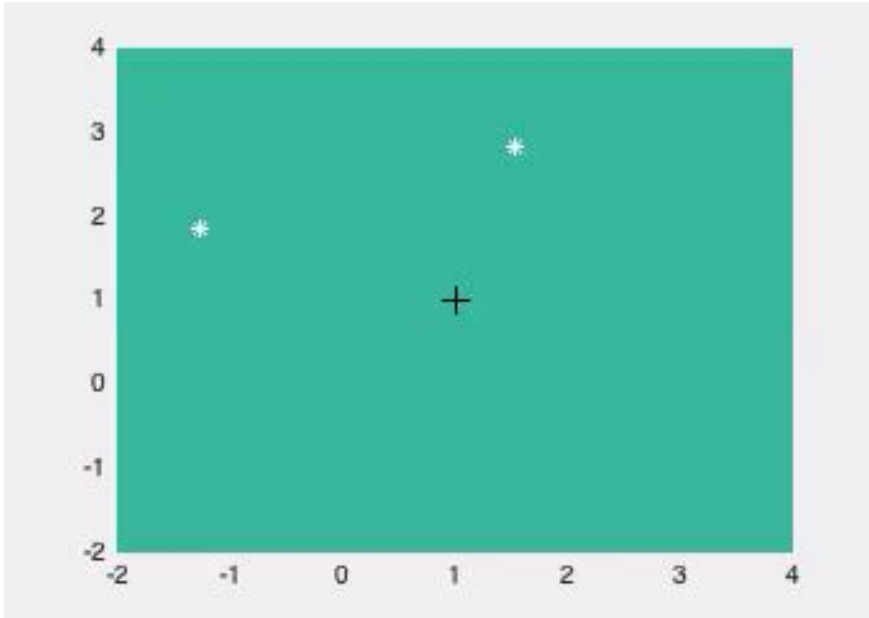
How much ensemble is sufficient?

- We just need more than N number of ensemble members if
 - we know the number of positive Lyapunov exponents
 - the system is well-observed
 - the model is linear
 - the model has no random errors
 - a deterministic Kalman filter (e.g., ETKF) is used
- Practically, we'd like to have as many ensemble members as possible

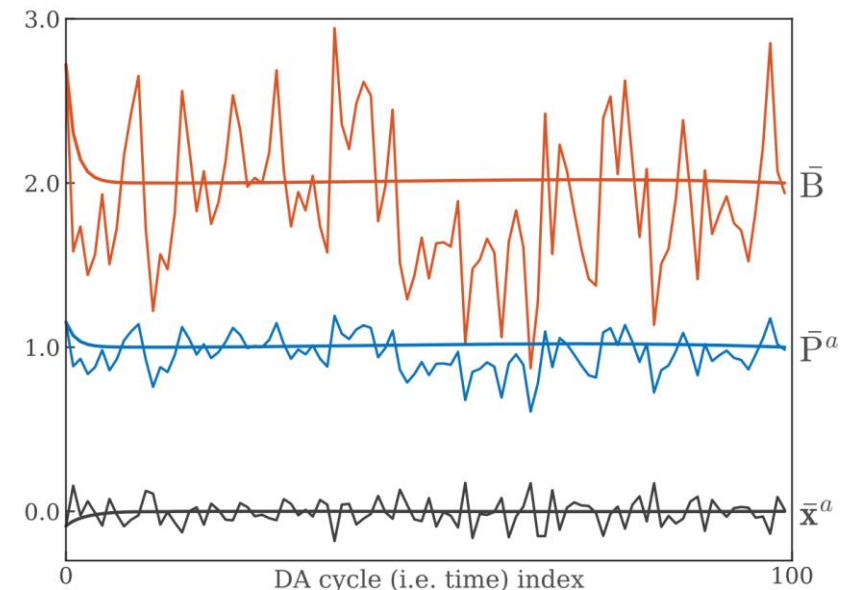


Sampling error

- The ensemble Kalman Filter theory assumes that the ensemble is large enough to give an accurate estimate of the sample mean and covariance, $\bar{\mathbf{x}}$ and \mathbf{P} .
- Sampling errors can be caused by:
 - Limited ensemble members due to constraints from computational resources
 - Model nonlinearity



Even for a two variable model a large sample size is needed to accurately estimate the mean and covariance: $\mathbf{x} \sim \mathcal{N}([1, 1], \mathbf{I}_2)$

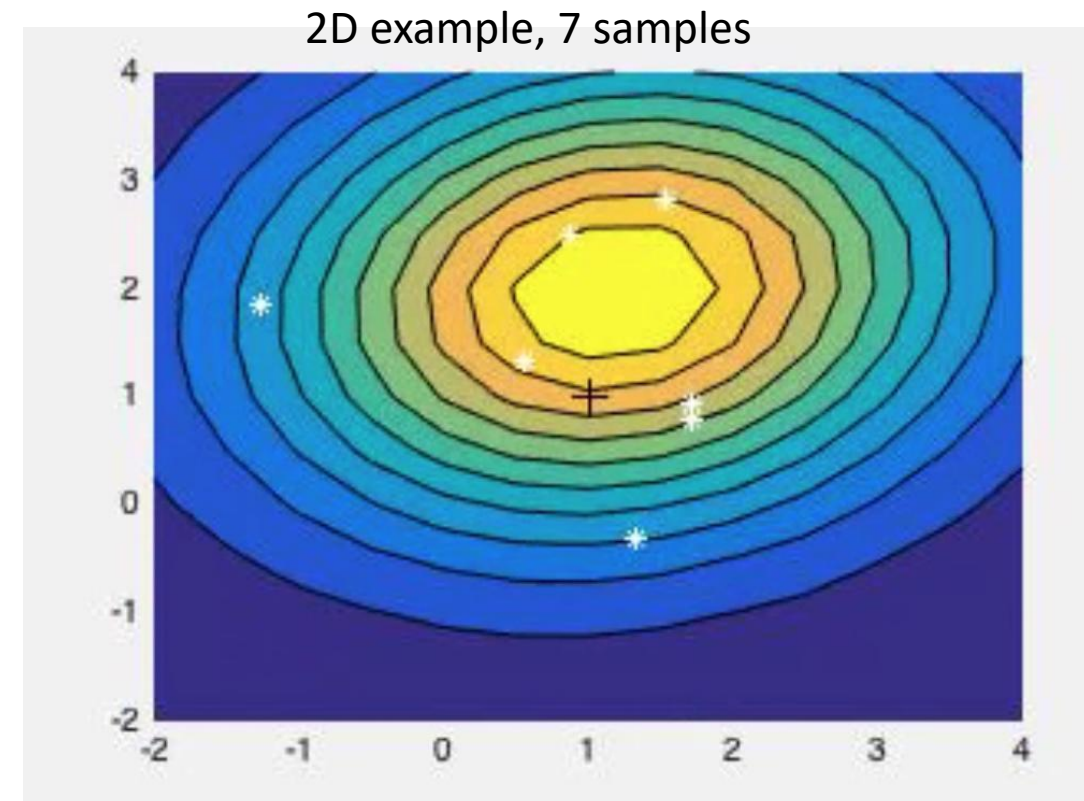


Sampling errors caused by non-linear model even if Gaussianity is preserved (jittery lines) compared to the linear model (smooth lines)

Consequences of under sampling in EnKF

There are numerous consequences to under sampling in the EnKF

1. There may be a **bias** in the **ensemble mean**.



Consequences of under sampling in EnKF

2. The forecast ensemble spread, which defines \mathbf{P}^f , will be subject to sampling error

- Recall

$$\mathbf{x}_k^{(i),a} = \mathbf{x}_k^{(i),f} + \mathbf{P}_k^f \mathbf{H}^T (\mathbf{H} \mathbf{P}_k^f \mathbf{H}^T + \mathbf{R})^{-1} (\mathbf{y}_k + \boldsymbol{\epsilon}_y^{(i)} - \mathbf{H} \mathbf{x}_k^{(i),f})$$

- If the spread (\mathbf{P}_k^f) is too large the analysis ensemble will over fit to the observations.
- If the spread (\mathbf{P}_k^f) is too small, the ensemble will under fit to the observations. If the ensemble repeatedly underestimates the forecast error and the information in the observations is ignored, then it is difficult to regain spread in the ensemble. This is called 'filter divergence'.

2D example, 7 samples

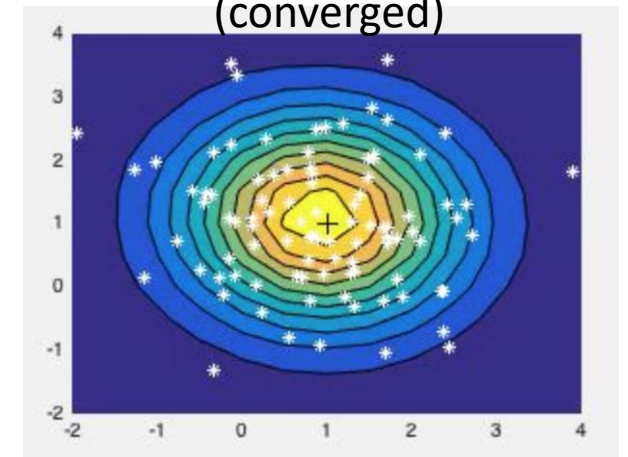
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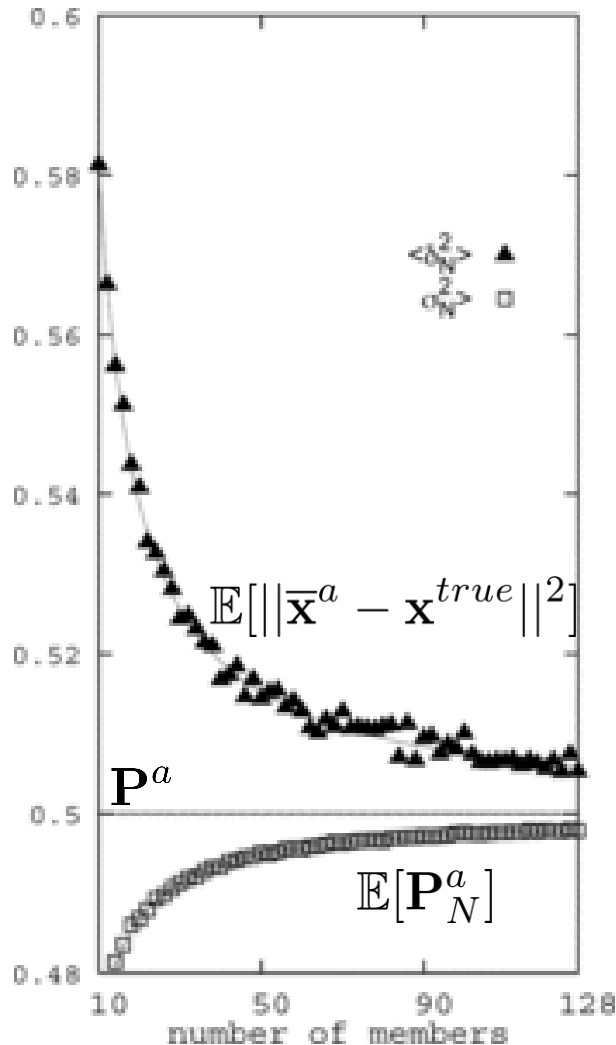
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2D example, 100 samples (converged)

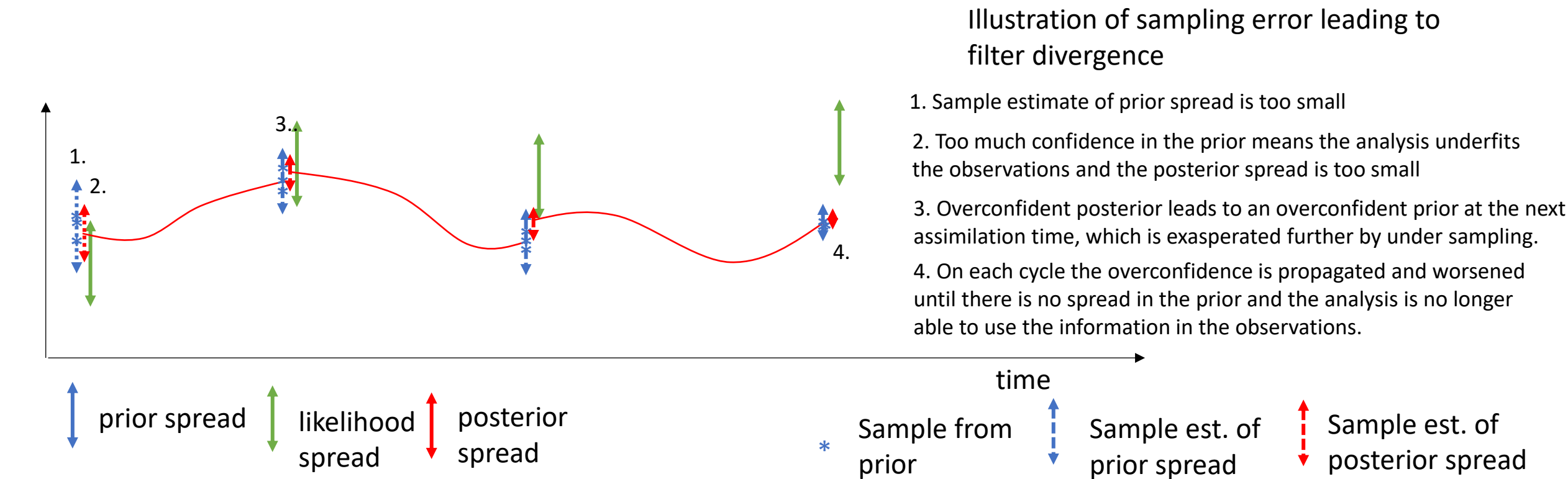
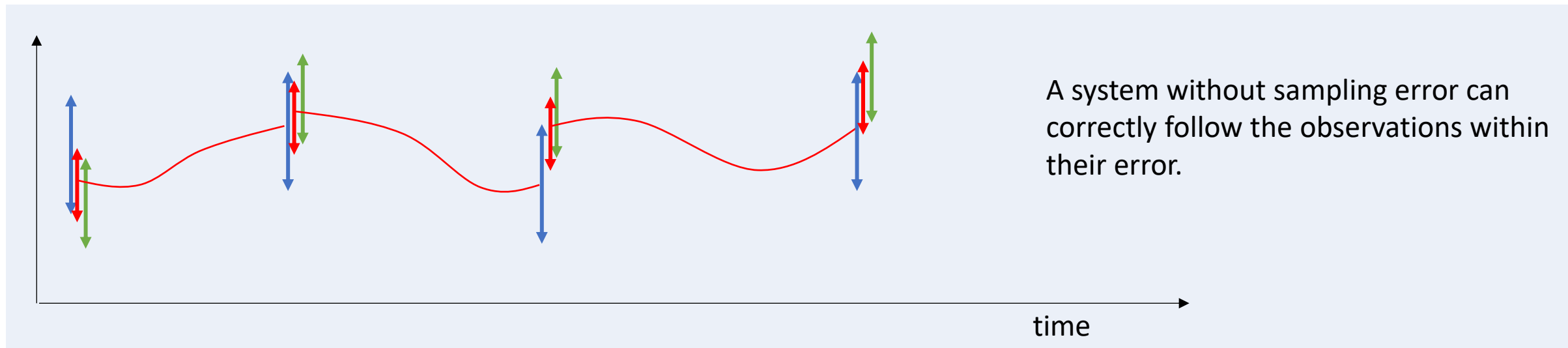


Variance bias is function of ensemble size



- In limit $N \rightarrow \infty$ the mean squared error and ensemble variance match on average.
- For small ensembles, the ensemble spread underestimates the analysis uncertainty.
- Smaller ensembles result in less error removal.

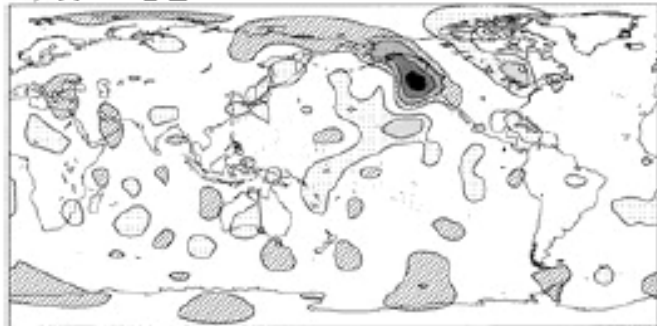
Image: Sacher, W., & Bartello, P. (2008). Sampling errors in ensemble Kalman filtering. Part I: Theory. *Monthly Weather Review*, 136(8), 3035-3049.



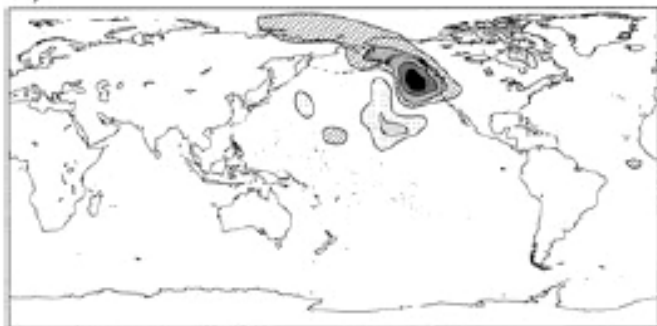
Consequences of under sampling in EnKF

3. The **correlation will be subject to sampling error**. Implying that observations can influence regions and variables that they shouldn't.

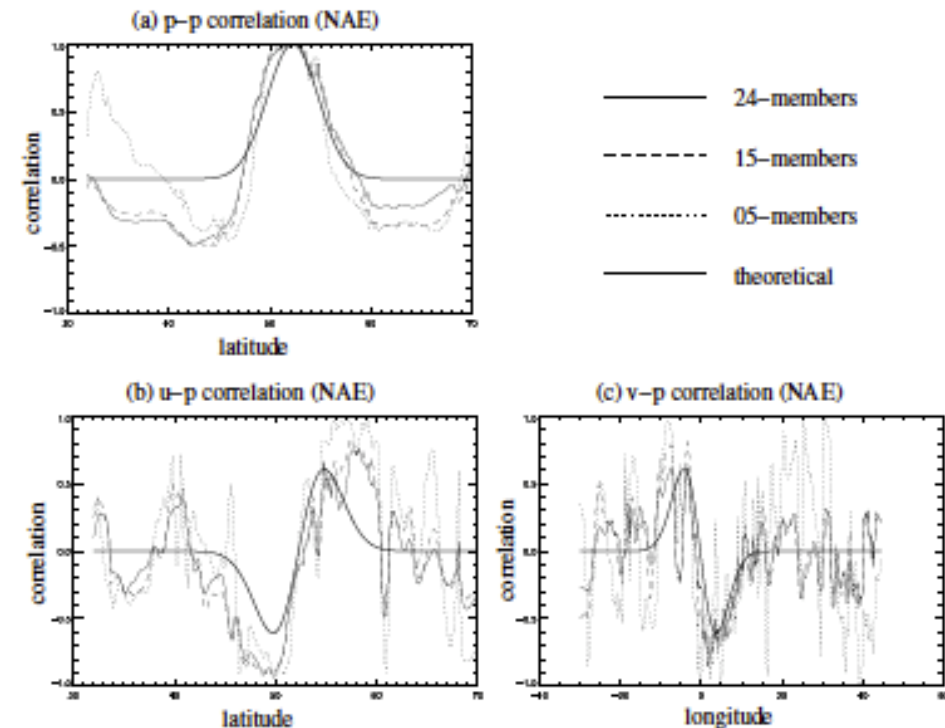
a) $N = 32$



c) $N = 128$



From Houtekamer & Mitchell (1998)



From Bannister, Migliorini & Dixon (2011)

Consequences of under sampling in EnKF

4. \mathbf{P}_k^f is rank deficient

Remember that $\mathbf{P}^f \in \mathbb{R}^{n \times n}$ and $\mathbf{X}'^f \in \mathbb{R}^{n \times N}$, so $\text{rank}(\mathbf{P}_k^f) = \text{rank}(\mathbf{X}_k'^f) \leq N-1 \ll n$.

The analysis increments are in the sub-space spanned by the forecast ensemble

- The analysis increments are given by

$$\mathbf{x}_k^{(i),a} - \mathbf{x}_k^{(i),f} = \mathbf{P}_k^f \mathbf{H}^T (\mathbf{H} \mathbf{P}_k^f \mathbf{H}^T + \mathbf{R})^{-1} (\mathbf{y}_k + \boldsymbol{\epsilon}_y^{(i)} - \mathbf{H} \mathbf{x}_k^{(i),f})$$

- The analysis increments are therefore a linear combination of the forecast error perturbations.
- Therefore, even if the observations indicate otherwise, the analysis is restricted to space spanned by the ensemble which has at the most a dimension of $N-1$.
- The error covariance matrix cannot represent all modes of forecast/analysis error.

Possible solutions

1. Use more ensemble members (see Miyoshi et al. 2014)
2. Re-centre the ensemble around a deterministic analysis e.g., from 4DVar.
 - Addresses problem of bias in the ensemble mean
3. Ensemble inflation
 - Addresses problem of filter divergence
4. Localization
 - Addresses problem of spurious correlations
 - Splits problem into quasi-independent problems, increasing the rank of forecast perturbation matrix.
5. Combine ensemble with variational approaches (see afternoon's lectures)
 - These are known as hybrid methods

Focus of this
lecture

Ensemble inflation



Ways to inflate

- Additive inflation (Mitchell and Houtekamer, 2000; Corazza *et al.*, 2003)
 - At each model time step add a random perturbation using similar ideas to representing model error given in the last lecture

$$\mathbf{x}_k^{(i)} = M_{t_{k-1} \rightarrow t_k}(\mathbf{x}_{k-1}^{(i)}) + \boldsymbol{\eta}_k^{(i)}, \text{ where } \boldsymbol{\eta} \sim N(\mathbf{0}, \mathbf{Q})$$

- Multiplicative inflation (Anderson and Anderson, 1999)

$$\mathbf{P}_{\text{inflated}}^f = (1 + \rho)^2 \mathbf{P}^f, \quad \rho > 0$$

- Relaxation to prior ensemble (Zhang *et al.*, 2004; Whitaker and Hamill, 2012)
 - Only accept part of the spread reduction proposed by

$$\mathbf{P}_k^a = (\mathbf{I} - \mathbf{K}_k \mathbf{H}) \mathbf{P}_k^f$$

- Two methods: relaxation to prior perturbation and relaxation to prior spread

Relaxation

- Relaxation to prior perturbation (Zhang & Sun, 2004):

$$\mathbf{x}'^{a,(n)} \leftarrow (1 - \alpha)\mathbf{x}'^{a,(n)} + \alpha\mathbf{x}'^{f,(n)} \text{ with } \mathbf{x}'^{(n)} = \mathbf{x}^{(n)} - \frac{1}{N} \sum_{m=1}^N \mathbf{x}^{(m)}$$

- Does not work with rotation in ETKF as there is not a one-to-one relation between forecast and analysis members.
- Relaxation to prior spread (Whitaker & Hamill, 2012):

$$\mathbf{x}_i'^{a,(n)} \leftarrow \mathbf{x}_i'^{a,(n)} \left(\alpha \frac{\sigma_i^b - \sigma_i^a}{\sigma_i^a} + 1 \right)$$

if $\alpha = 0$, $x_i'^{a,(n)} = x_i'^{a,(n)}$; if $\alpha = 1$, $x_i'^{a,(n)} = x_i'^{b,(n)}$

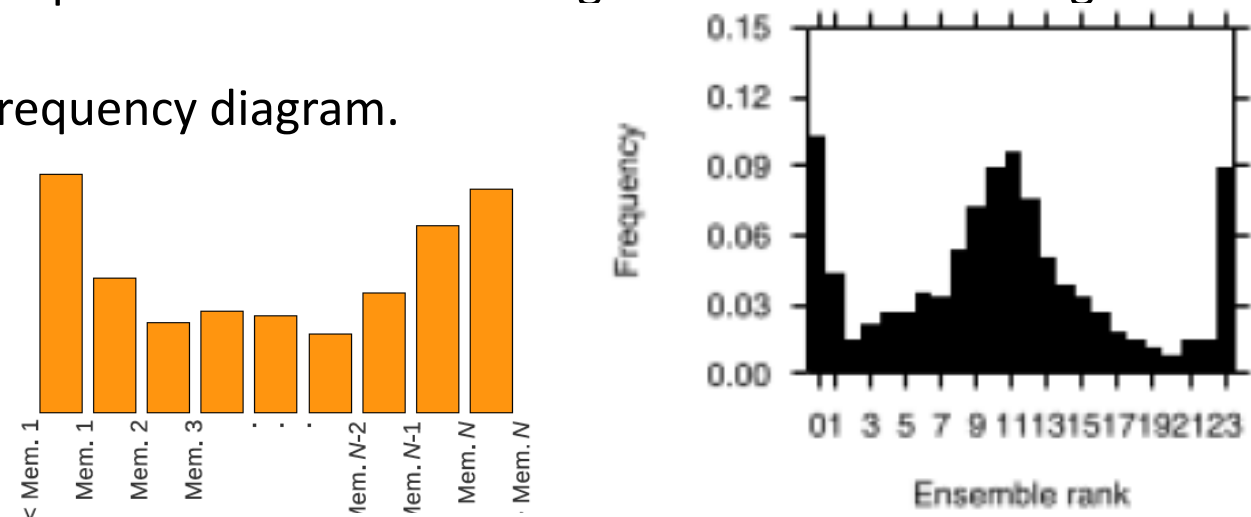
Tuning the inflation factor – validation of ensemble spread

Method 1: rank histograms (Hamill, T., 2001):

- For the ensemble to be reliable it is assumed that it is sampling the same distribution as the truth.
 - The ensemble and true distribution lead to the same number of samples for each interval of the PDF
- A rank histogram is constructed by considering a point in space that is well observed.
 - The values of the ensemble members at that point are ranked from highest to lowest creating $N-1$ bins.
 - Then each observation is binned to give a frequency diagram.

Interpretation:

- Concave shape- the ensemble is under spread
- Convex shaped- the ensemble is overspread
- Flat- the ensemble is correctly spread
- Asymmetric- the ensemble is biased



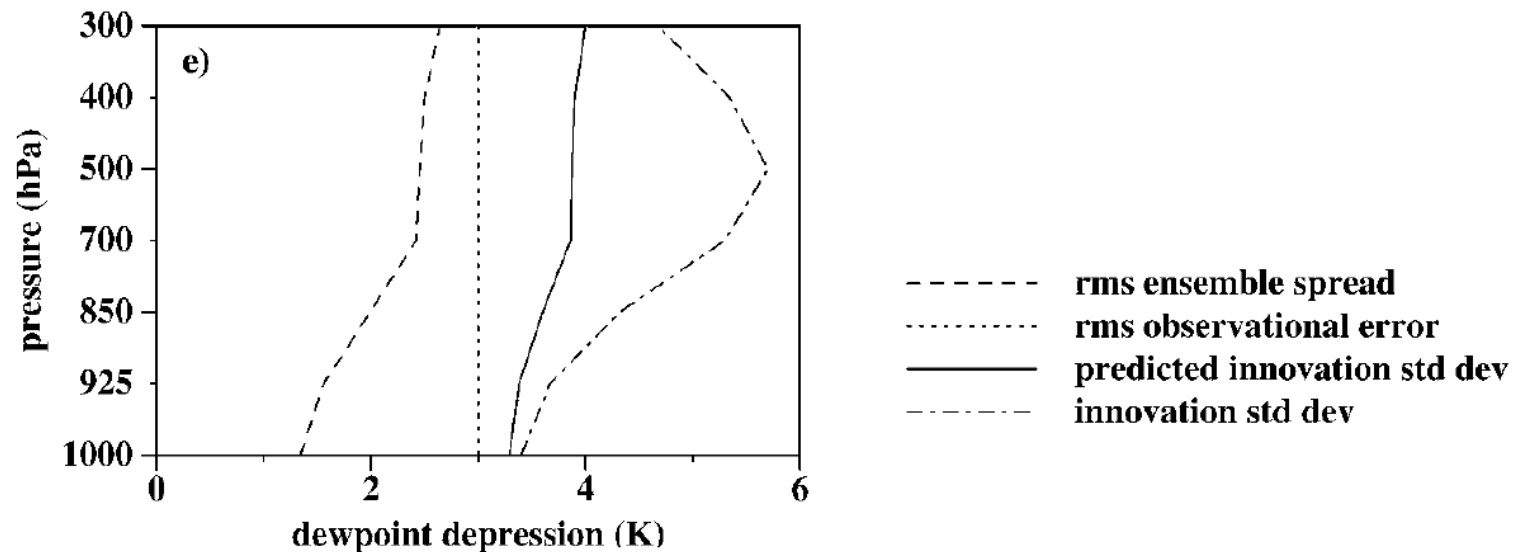
Rank histograms for surface precipitation rate rate. From [Migliorini et al. \(2011\)](#).

Tuning the inflation factor – validation of ensemble spread

Method 2: Covariance matching

- This checks that the sum of the spread in the background ensemble and observation error variance match with the variance of the innovations (e.g. Houtekamer et al. 2005).

$$E \left[(\mathbf{y} - \mathbf{H}\mathbf{x}^f)(\mathbf{y} - \mathbf{H}\mathbf{x}^f)^T \right] \approx \mathbf{R} + \mathbf{H}\mathbf{P}^f \mathbf{H}^T$$



The idea of covariance matching has led to various schemes for adaptive covariance inflation e.g. Kotsuki et al. 2017

FIG. 5. Comparison of error amplitudes that have been averaged over a 10-day experimental period. Shown are the predicted innovation std dev (solid) that should match the observed innovation std dev (dashed–dotted). The predicted std dev is computed from the rms observational error (dotted) and the rms ensemble spread (dashed).

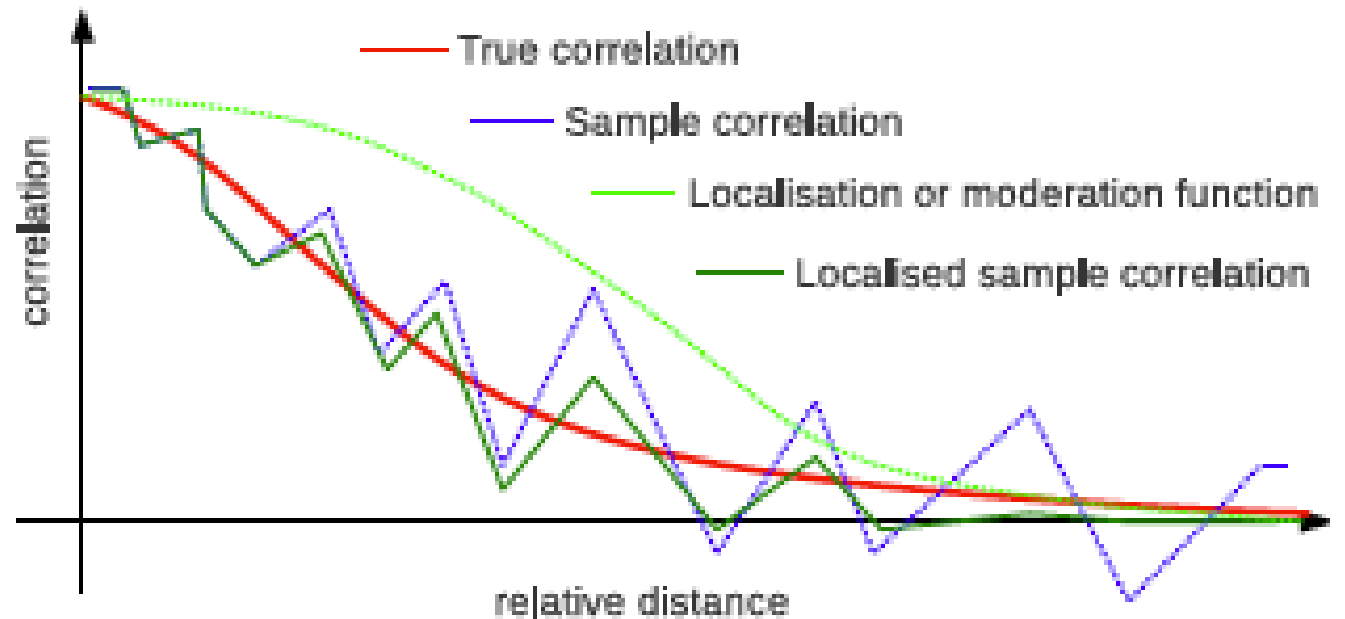
Localisation



The aim of localisation is to restrict the influence of observations to just a physically realistic region.

Two ways of doing this are:

- Covariance localisation (\mathbf{P}^f -localisation)
 - This modifies the forecast error covariance matrix to reduce long-range correlations.
- Domain localisation (\mathbf{R} -localisation)
 - This restricts observations which are allowed to influence each grid point.



Covariance localisation (Houtekamer and Mitchell, 2001)

- Instead of using directly, $\boldsymbol{\rho} \circ \mathbf{P}^f$ is used.
- In practice cannot act on \mathbf{P}^f directly

$$\mathbf{K} = \boldsymbol{\rho} \circ (\mathbf{P}^f \mathbf{H}^T) (\boldsymbol{\rho} \circ (\mathbf{H} \mathbf{P}^f \mathbf{H}^T) + \mathbf{R})^{-1}$$

- Need to choose function $\boldsymbol{\rho}$ and length scales, this may be state-dependent
- Not clear how to define distance between observations which have no clearly defined location in space, e.g. satellite observations
- Not clear how to deal with multivariate covariances
- Can affect the balance e.g. to conserve geostrophic balance length scale $O(1000)\text{km}$ must be used in the horizontal

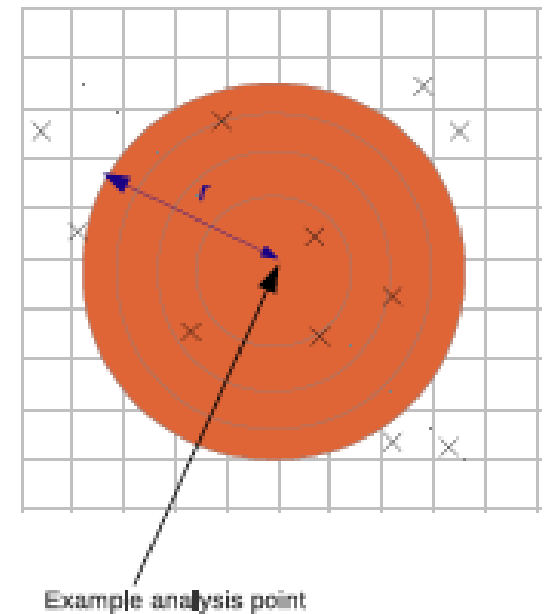
Domain localisation (Hunt et al. 2007)

- Each grid point performs an independent ETKF where only observations within a localisation radius are assimilated (LETKF)
- To avoid discontinuous DA corrections observational error variances are gradually inflated for observations far from selected grid point i , e.g.

$$f_{Rloc} = \exp \left[\frac{d(i,j)^2}{2L^2} \right]$$

where

- $d(i,j)$ is the distance between observation i and model grid point j .
- L is the localisation lengthscale.



Hybrid methods

Hybrid methods combine the best parts of the EnKF (flow-dependent \mathbf{P}^f) with the best parts of variational methods (full rank \mathbf{B}).

The earliest hybrid method was proposed by Hamil and Snyder (2000), in which the representation of the error covariance of the prior information is a weighted combination of the flow-dependent estimate from the EnKF, \mathbf{P}^f , and the full rank estimate used in variational methods \mathbf{P}^s

$$\mathbf{P}_{hybrid} = \beta \mathbf{P}^s + (1 - \beta) \mathbf{P}^f$$

where β is a tunable parameter.

Note localisation and inflation of the ensemble are still necessary.

Summary

- Ensemble data assimilation relies on a sample estimate of the mean and covariance of forecast distribution. This allows it to provide a flow-dependent estimate of the forecast uncertainty.
- If the ensemble size is much smaller than the size of the state then sampling error becomes an issue
 - Biases
 - Analysis increments lie in the subspace of the ensemble
 - Filter divergence
 - Spurious correlations
- To make ensemble DA practical need
 - Ensemble inflation
 - Localisation
 - ...Hybrid methods

Further reading

- Anderson JL, Anderson SL. 1999. A Monte Carlo implementation of the nonlinear filtering problem to produce ensemble assimilations and forecasts. *Mon. Weather Rev.* 127: 2741–2758.
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