

What Makes Filtering Hard?

Andrew M Stuart

Mathematics Institute, University of Warwick

cw/ S. Agapiou, O. Papaspiliopoulos, D. Sanz-Alonso
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cw/ D. Kelly

In Preparation

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Table of Contents

Filtering

Convergence of Filters

Linear Filtering Problem

Comparison of Optimal and Standard Proposals

Conclusions

Table of Contents

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Convergence of Filters

Linear Filtering Problem

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Problem Statement

Stochastic Dynamics Model

$$\begin{aligned}v_{j+1} &= \psi(v_j) + \xi_j, & \xi_j &\sim N(\mathbf{0}, \mathbf{Q}) \\y_{j+1} &= h(v_{j+1}) + \eta_{j+1}, & \eta_j &\sim N(\mathbf{0}, \mathbf{R}).\end{aligned}$$

Objective of Filtering

Define $\mu_0 = \mathbb{P}(v_0)$, $\mu_j = \mathbb{P}(v_j | Y_j)$, $Y_j := (y_1, \dots, y_j)$.

Sequentially Update $\mu_j \mapsto \mu_{j+1}$.

Propagation of Probabilities I

Using

$$\mathbb{P}(v_{j+1}|Y_j) = \mathbb{P}(v_{j+1}|v_j)\mathbb{P}(v_j|Y_j)$$

we obtain

Densities

$$\mathbb{P}(v_{j+1}|Y_{j+1}) = \frac{1}{\mathbb{P}(y_{j+1}|Y_j)} \mathbb{P}(y_{j+1}|v_{j+1}) \int \mathbb{P}(v_{j+1}|v_j)\mathbb{P}(v_j|Y_j)dv_j.$$

Measures

$$\mu_{j+1} = L_j P \mu_j.$$

Propagation of Probabilities II

Using

$$\mathbb{P}(v_{j+1}|v_j, y_{j+1}) = \frac{1}{\mathbb{P}(y_{j+1}|v_j)} \mathbb{P}(y_{j+1}|v_{j+1}) \mathbb{P}(v_{j+1}|v_j)$$

we also find that

Densities

$$\mathbb{P}(v_{j+1}|Y_{j+1}) = \frac{1}{\mathbb{P}(y_{j+1}|Y_j)} \int \mathbb{P}(v_{j+1}|v_j, y_{j+1}) \mathbb{P}(y_{j+1}|v_j) \mathbb{P}(v_j|Y_j) dv_j.$$

Measures

$$\mu_{j+1} = Q_j \mu_j.$$

Sampling and Particle Filters

Monte Carlo Sampling

$$S^N \mu(du) = \sum_{i=1}^N \frac{1}{N} \delta_{u^i}(du),$$
$$u^i \sim \mu.$$

Standard Proposal

$$\mu_{j+1}^N = L_j S^N P \mu_j^N, \quad \mu_0^N = \mu_0.$$

Optimal Proposal

$$\mu_{j+1}^N = S^N Q_j \mu_j^N, \quad \mu_0^N = S^N \mu_0.$$

The Standard Proposal

$$\mu_0^N = \mu_0; \quad \mu_j^N(du) = \sum_{i=1}^N w_j^i \delta_{u_j^i}(du), \quad j \geq 1,$$

$$u_{j+1}^i = \psi(u^i) + \xi_j^i, \quad u_j^i \sim \mu_j^N, \quad \xi_j^i \sim N(0, Q),$$

$$\widehat{w}_{j+1}^i = \exp\left(-\frac{1}{2}|y_{j+1} - h(u_{j+1}^i)|_{\mathbf{R}}^2\right),$$

$$w_{j+1}^i = \widehat{w}_{j+1}^i / \sum_{m=1}^N \widehat{w}_{j+1}^m.$$

The Optimal Proposal (case $h(\cdot) = H\cdot$)

$$\mu_j^N(du) = \sum_{i=1}^N \frac{1}{N} \delta_{u_j^i}(du), \quad j \geq 0,$$

$$\hat{u}_{j+1}^i = (I - KH)\psi(u_j^i) + Ky_{j+1} + \zeta_j^i, \quad \zeta_j^i \sim N(0, C)$$

$$\hat{w}_{j+1}^i = \exp\left(-\frac{1}{2}|y_{j+1} - H\psi(u_j^i)|_S^2\right)$$

$$w_{j+1}^i = \hat{w}_{j+1}^i / \sum_{m=1}^N \hat{w}_{j+1}^m$$

$$\mu_{j+1}^N(du) = S^N \left(\sum_{i=1}^N w_{j+1}^i \delta_{\hat{u}_{j+1}^i}(du) \right).$$

Here

$$S = HQH^* + R, \quad K = QH^*S^{-1},$$

$$C = (I - KH)Q.$$

EnKF (case $h(\cdot) = H\cdot$)

$$\mu_j^N(du) = \sum_{i=1}^N \frac{1}{N} \delta_{u_j^i}(du), \quad j \geq 0,$$

$$u_{j+1}^i = (I - K_{j+1}H)\psi(u_j^i) + K_{j+1}y_{j+1} + \zeta_j^i, \quad \zeta_j^i \sim N(0, C_{j+1})$$

Here \widehat{C}_{j+1} is the empirical covariance of the predicted particle positions

$$\widehat{u}_{j+1}^i = \psi(u_j^i) + \zeta_j^i$$

and

$$S_{j+1} = H\widehat{C}_{j+1}H^* + R, \quad K_{j+1} = \widehat{C}_{j+1}H^*S^{-1}, \\ C_{j+1} = (I - K_{j+1}H)Q.$$

References

It's all about the weights



T. Bengtsson, P. Bickel, B. Li, *Curse-of-dimensionality revisited: Collapse of the particle filter in very large scale systems*, Probability and Statistics: Essays in Honour of D.A. Freeman, 2008.



C. Snyder, T. Bengtsson, P. Bickel, *Obstacles to high-dimensional particle filtering*, Monthly Weather Review, **136**(2008).



C. Snyder, *Particle filters, the 'optimal' proposal and high-dimensional systems*, In Proceedings of the ECMWF Seminar on Data Assimilation for Atmosphere and Ocean, 2011.



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Table of Contents

Filtering

Convergence of Filters

Linear Filtering Problem

Comparison of Optimal and Standard Proposals

Conclusions

Metric on Random Measures

Let $(\mathcal{X}, \mathcal{F})$ be an arbitrary measurable space and ν a probability measure. Then

$$\nu(\phi) := \int_{\mathcal{X}} \phi(u) \nu(du).$$

If ν_1, ν_2 are random probability measures then we introduce a metric:

$$d(\nu_1, \nu_2) := \left(\sup_{|\phi| \leq 1} \mathbb{E} \left[(\nu_1(\phi) - \nu_2(\phi))^2 \right] \right)^{\frac{1}{2}}.$$

P. Rebeschini and R. Van Handel. Can local particle filters beat the curse of dimensionality? *Ann. Appl. Prob.* **25**(2015).

P. Del Moral. Feynman-Kac formulae: genealogical and interacting particle systems with applications 2004.

The Standard Proposal

Recall

$$\begin{aligned}\mu_{j+1} &= L_j P \mu_j, \\ \mu_{j+1}^N &= L_j \mathbf{S}^N P \mu_j^N, \mu_0^N = \mu_0.\end{aligned}$$

Theorem 1

Assume that h is a bounded nonlinear function. Then there exists $\kappa = \kappa(Y_J)$ such that

$$d(\mu_j, \mu_j^N)^2 \leq \frac{\kappa}{N}.$$

P. Rebeschini and R. Van Handel. Can local particle filters beat the curse of dimensionality? Ann. Appl. Prob. **25**(2015).

$P\mu_j$ is proposal for **importance sampling** of μ_{j+1} .

The Optimal Proposal

Recall

$$\begin{aligned}\mu_{j+1} &= Q_j \mu_j, \\ \mu_{j+1}^N &= S^N Q_j \mu_j^N, \mu_0^N = S^N \mu_0.\end{aligned}$$

Theorem 2

Assume that ψ is a bounded nonlinear function and $h(\cdot) = H \cdot$ is linear. Then there exists $\kappa = \kappa(Y_J)$ such that

$$d(\mu_j, \mu_j^N)^2 \leq \frac{\kappa}{N}.$$

D.T.B. Kelly and A.M. Stuart. In preparation (2016).

Q_j factors into **importance sampling** followed by conditioned dynamics when applied to a sum of Dirac measures.

The EnKF

Theorem 3

Assume that $\psi(\cdot) = M\cdot$ and $h(\cdot) = H\cdot$ is linear. Then there exists $\kappa = \kappa(Y_J)$ such that

$$d(\mu_j, \mu_j^N)^2 \leq \frac{\kappa}{N}.$$

F. Le Gland, V. Monbet and V.-D. Tran. INRIA Report 7014(2009).

The paper also identifies the limit of μ_j^N obtained in the nonlinear, non-Gaussian setting and shows that it is not μ_j . Suggests using EnKF as a proposal for **importance sampling**.

Autonormalized Importance Sampling (IS)

π and μ probability measures in the arbitrary measurable space $(\mathcal{X}, \mathcal{F})$. Assume $\mu \ll \pi$ so that

$$\mu(du) = g(u)\pi(du) / \int_{\mathcal{X}} g(u)\pi(du).$$

Applying the Monte Carlo sampling operator twice gives

$$\mu(\phi) = \frac{\pi(\phi g)}{\pi(g)} \approx \frac{S^N \pi(\phi g)}{S^N \pi(g)} := \mu^N(\phi)$$

IS Convergence

Thus we have the approximation of μ defined by, for $u^i \sim \pi$,

$$\mu^N = \sum_{i=1}^N w^i \delta_{u^i}, \quad w^i = \frac{g(u^i)}{\sum_{k=1}^N g(u^k)}.$$

Define $\rho := \frac{\pi(g^2)}{\pi(g)^2} \in [1, \infty]$.

Theorem 4

$$d(\mu^N, \mu)^2 \leq \frac{4}{N} \rho.$$

Thus ρ , and its dependence on the problem at hand, is key.

Table of Contents

Filtering

Convergence of Filters

Linear Filtering Problem

Comparison of Optimal and Standard Proposals

Conclusions

IS for Filtering

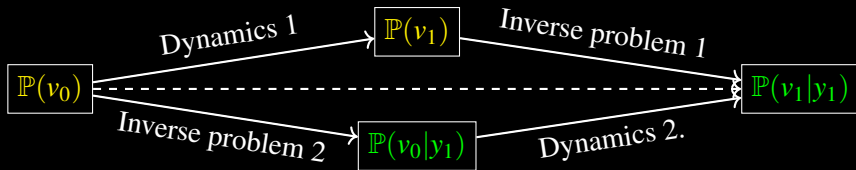
Filtering

Signal: $v_1 = Mv_0 + N(0, Q), \quad v_0 \sim N(0, P) = \mathbb{P}_0.$

Data: $y_1 = Hv_1 + N(0, R).$ Find: $\mathbb{P}(v_1|y_1).$

Standard proposal: Top Route: dynamics, then inverse problem.

Optimal proposal: Bottom Route: inverse problem, then dynamics.



Inverse Problem: Standard Proposal

Filtering

Signal: $v_1 = Mv_0 + N(0, Q), \quad v_0 \sim N(0, P) = \mathbb{P}_0.$

Data: $y_1 = Hv_1 + N(0, R).$ Target: $\mu := \mathbb{P}(v_1|y_1).$

Behaviour of filtering model determined by inverse problem

$$y_1 = Ku + \eta, \quad u \sim N(0, \Sigma), \quad \eta \sim N(0, \Gamma),$$

with

- Target $\mu = \mathbb{P}(v_1|y_1) : K = H, \Gamma = R, \Sigma = MPM^* + Q$
- Proposal $\pi_{st} = \mathbb{P}(v_1).$

Inverse Problem: Optimal Proposal

Filtering

Signal: $v_1 = Mv_0 + N(0, Q), \quad v_0 \sim N(0, P) = \mathbb{P}_0.$

Data: $y_1 = Hv_1 + N(0, R).$ Target: $\mu := \mathbb{P}(v_0|y_1).$

Behaviour of filtering model determined by inverse problem

$$y_1 = Ku + \eta, \quad u \sim N(0, \Sigma), \quad \eta \sim N(0, \Gamma),$$

with

- Target $\mu = \mathbb{P}(v_0|y_1) : K = HM, \Gamma = HQH^* + R, \Sigma = Q$
- Proposal $\pi_{st} = \mathbb{P}(v_0).$

Linear Bayesian Inverse Problems in Hilbert Space

Interested in recovering $u \in \mathcal{H}$ from data $y = Ku + \eta \in \mathcal{H}$.

Bayesian Inverse Problem

$$\left. \begin{array}{l} \text{Prior (proposal): } u \sim \pi = N(0, \Sigma) \\ \text{Data: } y = Ku + \eta \in \mathcal{Y}, \quad \eta \sim N(0, \Gamma) \end{array} \right\} u|y \sim \mu = N(m, C).$$

Linear Bayesian Inverse Problems in Hilbert Space

Interested in recovering $u \in \mathcal{H}$ from data $y = Ku + \eta \in \mathcal{H}$.

Bayesian Inverse Problem

$$\left. \begin{array}{l} \text{Prior (proposal)} : u \sim \pi = N(0, \Sigma) \\ \text{Data} : y = Ku + \eta \in \mathcal{Y}, \quad \eta \sim N(0, \Gamma) \end{array} \right\} u|y \sim \mu = N(m, C).$$

Let $y_0 = \Gamma^{-\frac{1}{2}}y$ and $S = \Gamma^{-\frac{1}{2}}K\Sigma^{\frac{1}{2}}$. Then

Equivalent Bayesian Inverse Problem

$$\left. \begin{array}{l} \text{Prior (proposal)} : u_0 \sim \pi_0 = N(0, I) \\ \text{Data} : y_0 = Su_0 + \eta_0 \in \mathcal{Y}_0, \quad \eta_0 \sim N(0, I) \end{array} \right\} u_0|y_0 \sim \mu_0 = N(m_0, C_0).$$

Intrinsic Dimension

Key: Eigenvalues of $A := S^*S = \Sigma^{1/2}K^*\Gamma^{-1}K\Sigma^{1/2}$.

$$\text{efd} = \text{Tr} \left((I + A)^{-1}A \right)$$

$$\tau = \text{Tr} (A)$$

(efd) Spiegelhalter, Best, Carlin, Van der Linde, Bayesian measures of model complexity and fit, 2002.

(τ) Bengtsson, Bickel, and Li, Curse-of-dimensionality revisited: Collapse of the particle filter in very large systems, 2008.

Lemma 5

$$\frac{1}{\|I + A\|} \tau \leq \text{efd} \leq \tau.$$

$$\text{efd} = \text{Tr} \left((\Sigma - C)\Sigma^{-1} \right) \leq \min\{d_u, d_y\},$$

$$\tau = \text{Tr} \left((\Sigma - C)C^{-1} \right).$$

Link Between τ , efd, ρ and $\mu \ll \pi$

Theorem 6

Let $\nu(du, dy) = \mathbb{P}(dy|u)\pi(du)$ and assume A bdd. Equivalent:

i) $\text{efd} < \infty$.

ii) $\tau < \infty$.

iii) For ν -a.a. y , $\mu(du) \propto g(u; y)\pi(du)$ and

$$0 < \pi(g(\cdot; y)) < \infty.$$

iv) It holds $0 < g(u; y) < \infty$ ν -a.s. and for ν -a.a. y

$$\rho := \frac{\pi(g(\cdot; y)^2)}{\pi(g(\cdot; y))^2} < \infty.$$

Link Between τ , efd, ρ and $\mu \ll \pi$

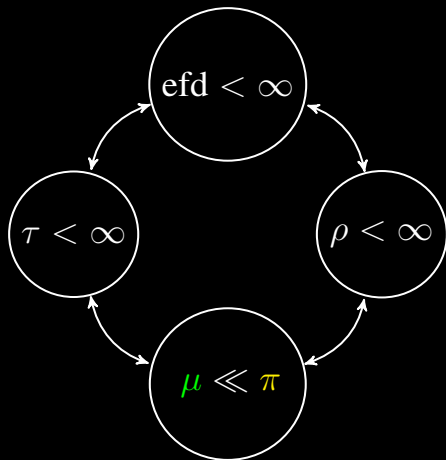


Table of Contents

Filtering

Convergence of Filters

Linear Filtering Problem

Comparison of Optimal and Standard Proposals

Conclusions

Ordering by τ

Filtering

Signal: $v_1 = Mv_0 + N(0, Q), \quad v_0 \sim N(0, P) = \mathbb{P}_0.$

Data: $y_1 = Hv_1 + N(0, R).$ Target: $\mathbb{P}(u|y_1), u = (v_0, v_1).$

	Standard Proposal	Optimal Proposal
S	$R^{-\frac{1}{2}}H(MPM^* + Q)^{\frac{1}{2}}$	$(HQH^* + R)^{-\frac{1}{2}}HMP^{\frac{1}{2}}$

Recall that $A = S^*S$, $\text{efd} = \text{Tr} \left((I + A)^{-1}A \right), \tau = \text{Tr} (A).$

Theorem 7

$$\tau_{st} \geq \tau_{op}$$

Diagonal Example: Small Noise and High Dimensions

$M = H = I \in \mathbb{R}^{d \times d}$, and $R = rI$, $Q = qI$, with $r, q > 0$.

Stationary regime $P_\infty = \frac{\sqrt{q^2 + 4qr} - q}{2} I$,

Chorin and Morzfeld. Conditions for successful data assimilation, 2013.

$$A_{st} = \frac{\sqrt{q^2 + 4qr} + q}{2r} I, \quad A_{op} = \frac{\sqrt{q^2 + 4qr} - q}{2(q+r)} I.$$

Outside stationarity: $P = pI$, with $p > 0$ then

$$A_{st} = \frac{p+q}{r} I, \quad A_{op} = \frac{p}{q+r} I.$$

Diagonal Example: Small Noise and High Dimensions

Regime	Param.	$\text{eig}(A_{st})$	$\text{eig}(A_{op})$	ρ_{st}	ρ_{op}
Noise	$r \rightarrow 0$	\mathbf{r}^{-1}	r	$\mathcal{O}(\mathbf{r}^{-d/2})$	$\mathcal{O}(1)$
	$r = q \rightarrow 0$	1	1	$\mathcal{O}(1)$	$\mathcal{O}(1)$
Large d	$d \rightarrow \infty$	1	1	$\mathcal{O}(\exp(\mathbf{d}))$	$\mathcal{O}(\exp(\mathbf{d}))$

Table: $P = P_\infty$.

Regime	Param.	$\text{eig}(A_{st})$	$\text{eig}(A_{op})$	ρ_{st}	ρ_{op}
Noise	$r \rightarrow 0$	\mathbf{r}^{-1}	1	$\mathcal{O}(\mathbf{r}^{-d/2})$	$\mathcal{O}(1)$
	$r = q \rightarrow 0$	\mathbf{r}^{-1}	\mathbf{r}^{-1}	$\mathcal{O}(\mathbf{r}^{-d/2})$	$\mathcal{O}(\mathbf{r}^{-d/2})$
Large d	$d \rightarrow \infty$	1	1	$\mathcal{O}(\exp(\mathbf{d}))$	$\mathcal{O}(\exp(\mathbf{d}))$

Table: $P = pI$.

Table of Contents

Filtering

Convergence of Filters

Linear Filtering Problem

Comparison of Optimal and Standard Proposals

Conclusions

Highlights

● GENERAL FRAMEWORK

- Standard and optimal filters converge in the large N limit.
- Cost determined by embedded importance sampling problem.

● INVERSE PROBLEMS

- Established connection between notions of dimension: efd and τ .
- Linked intrinsic dimensions to importance sampling through ρ .
- Link with absolute continuity in Hilbert space setting.

● FILTERING

- Theory from inverse problems determines proposal performance.
- Optimal proposal has lower dimension than standard proposal.
- Effect of dimension and small noise can be studied.