



A posteriori diagnostics

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Outline

1. Introduction
2. « J_{\min} » diagnostics
3. Observation space diagnostics
4. Combination of diagnostics
5. Conclusion

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Introduction

« A posteriori diagnostics »: diagnostics made after the analysis.

A noncomprehensive list of references on the subject!

- *Bennett 1992*: J_{\min} , ...
- *Talagrand 2000, 2010, 2015*: « a posteriori diagnostics », J_{\min} , DFS, ...
- *Desroziers and Ivanov 2001, Chapnik et al 2006, Desroziers et al 2009*,
- *Michel 2014*: « a posteriori diagnostics », J_{\min} , DFS, σ^o , σ^b tuning ...
- *Fisher 2003*: DFS, ...
- *Desroziers et al 2005*: diagnostics in observation space ...
- *Todling 2015, Howes 2015*: diagnostic of model error, ...
- *Waller et al 2015*: study of observation space diagnostics ...
- *Ménard 2000, 2015a, 2015b*: study of « a posteriori diagnostics » ...
- ...

General framework

- Statistical linear estimation

$$\mathbf{x}^a = \mathbf{x}^b + \delta\mathbf{x} = \mathbf{x}^b + \mathbf{K} \mathbf{d} = \mathbf{x}^b + \mathbf{B}\mathbf{H}^T (\mathbf{H}\mathbf{B}\mathbf{H}^T + \mathbf{R})^{-1} \mathbf{d},$$

with $\mathbf{d} = \mathbf{y}^o - \mathbf{H}(\mathbf{x}^b)$, innovation, \mathbf{K} , gain matrix,

\mathbf{B} et \mathbf{R} , covariances of background and observation errors.

- Also solution of the variational problem

$$J(\delta\mathbf{x}) = \delta\mathbf{x}^T \mathbf{B}^{-1} \delta\mathbf{x} + (\mathbf{d} - \mathbf{H} \delta\mathbf{x})^T \mathbf{R}^{-1} (\mathbf{d} - \mathbf{H} \delta\mathbf{x}).$$

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J_{\min} / χ^2

- Value of the cost function at its minimum

$$\begin{aligned} J_{\min} &= J^b(\delta \mathbf{x}^a) + J^o(\delta \mathbf{x}^a) \\ &= (\mathbf{Kd})^T \mathbf{B}^{-1} (\mathbf{Kd}) + ((\mathbf{I}-\mathbf{HK})\mathbf{d})^T \mathbf{R}^{-1} ((\mathbf{I}-\mathbf{HK})\mathbf{d}) \\ &= \mathbf{d}^T (\mathbf{HBH}^T + \mathbf{R})^{-1} \mathbf{HBH}^T (\mathbf{HBH}^T + \mathbf{R})^{-1} \mathbf{d} + \mathbf{d}^T (\mathbf{HBH}^T + \mathbf{R})^{-1} \mathbf{R} (\mathbf{HBH}^T + \mathbf{R})^{-1} \mathbf{d} \\ &= \mathbf{d}^T (\mathbf{HBH}^T + \mathbf{R})^{-1} \mathbf{d} \\ &= \mathbf{d}^T \mathbf{D}^{-1} \mathbf{d}. \end{aligned}$$

- Statistical expectation of J_{\min}

$$\begin{aligned} E(J_{\min}) &= E(\mathbf{d}^T \mathbf{D}^{-1} \mathbf{d}) \\ &= \text{Tr}(\mathbf{D}^{-1} E(\mathbf{d} \mathbf{d}^T)) \\ &= p, \text{ number of observations, if } \mathbf{D} = E(\mathbf{d} \mathbf{d}^T). \end{aligned}$$

- Unnecessary: $\mathbf{HBH}^T = \alpha \mathbf{HB}^t \mathbf{H}^T$ and $\mathbf{R} = \alpha \mathbf{R}^t$, $\delta \mathbf{x}^a$ optimal but $E(J_{\min}) = p / \alpha$!
- Statistical distribution of J_{\min}
- ✓ Gaussian \mathbf{d} : J_{\min} follows a χ^2 distribution of order p , with mean p and var. $2p$.
- ✓ Central limit theorem: J_{\min} follows a $\mathcal{N}(p, 2p)$ law, if p is large.

Expectation of subparts of J_{\min}

- J^b part

$$E(J_{\min}^b) = \text{Tr}(\mathbf{H}\mathbf{K})$$

= total DFS (Degrees of freedom for Signal)
= number of degrees of freedom,
provided by observations, to deviate from background.

- J^o part

$$E(J_{\min}^o) = p - \text{Tr}(\mathbf{H}\mathbf{K})$$

= total DFN (Degrees of freedom for Noise)
= number of degrees of freedom ,
provided by background, to deviate from observations.

- J^o subparts

$$E(J_{i, \min}^o) = p_i - \text{Tr}(\mathbf{\Pi}_i \mathbf{H}\mathbf{K} \mathbf{\Pi}_i^T),$$

$\mathbf{\Pi}_i$ selection of subset of observations i , with number p_i .

- $\text{Tr}(\mathbf{\Pi}_i \mathbf{H}\mathbf{K} \mathbf{\Pi}_i^T) = \text{DFS}_i$
= partial number of degrees of freedom,
provided by observations i , to deviate from background.

Computation of subparts of \mathbf{J}_{\min} in a variational assimilation

- Matrix \mathbf{K} is not explicitly computed in a variational assimilation.
- Computation of $\text{Tr}(\Pi_i \mathbf{H} \mathbf{K} \Pi_i^T)$ with a Monte Carlo procedure

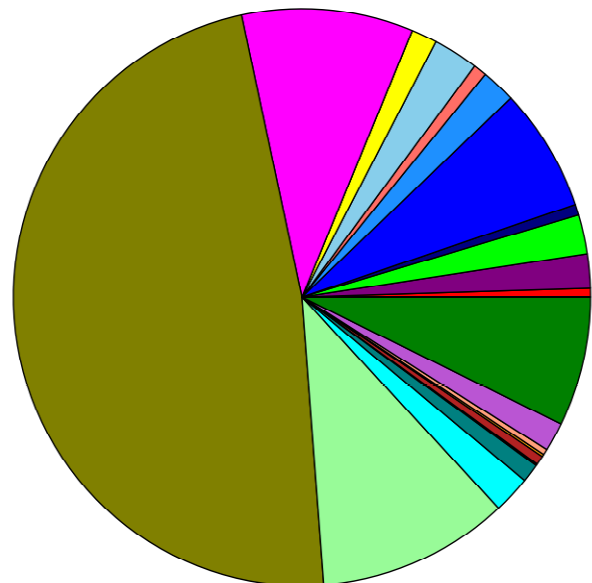
$$\begin{aligned}\text{Tr}(\Pi_i \mathbf{H} \mathbf{K} \Pi_i^T) &= \text{Tr}(\Pi_i^T \Pi_i \mathbf{H} \mathbf{K}) \\ &= \mathbb{E}(\delta \mathbf{y}_i^o{}^T \mathbf{R}_i^{-1} \mathbf{H}_i \mathbf{K} \delta \mathbf{y}_i^o) \\ &= \mathbb{E}(\delta \mathbf{y}_i^o{}^T \mathbf{R}_i^{-1} \mathbf{H}_i \delta \mathbf{x}^a(\delta \mathbf{y}_i^o)),\end{aligned}$$

with $\delta \mathbf{y}_i^o = \mathcal{N}(\mathbf{0}, \mathbf{R})$.

(Desroziers and Ivanov 2001, Desroziers et al 2009, Michel 2014)

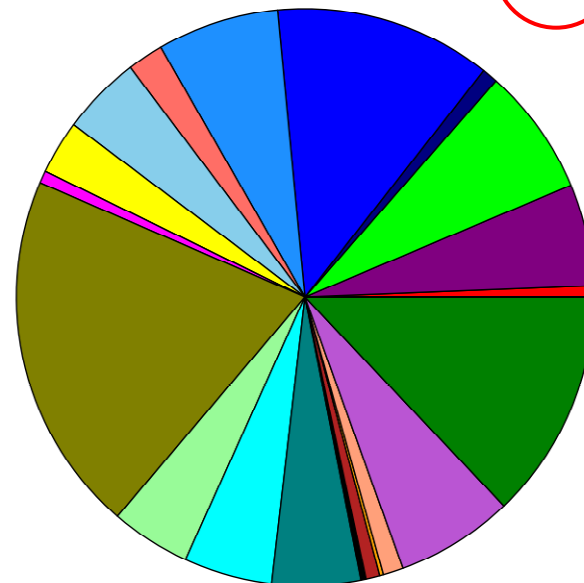
DFS_i computation

Proportions des nombres d'observations utilisées par type d'obs
 analyses cut-off long - ARPEGE metropole dbl
 observations conventionnelles et satellites
 cumul du nombre d'observations utilisées sur la période 2015120100 - 2015120118 : 20675304



GPS ground	0.49%	SSMIS	1.44%	SYNOP/SYNOR/RADNMF	0.50%
GPS sat	1.91%	GMI	0.00%	SHIP	0.14%
SATOB	2.20%	AIRS	9.60%	PILOT/PRF	0.32%
ATOVS HTRS	0.60%	TAST	47.84%	TFMP	1.61%
ATOVS AMSU-A	6.93%	CRIS	10.73%	AIRCRAFTS	7.25%
ATOVS AMSU-B	1.92%	GEORAD	2.10%	RADAR Vr	0.00%
SAPHIR	0.74%	SCATT	1.04%	RADAR Hvr	0.00%
ATMS	2.53%	BUOY	0.10%	BOGUS	0.00%

Part des DFS par type d'obs
 analyses cut-off long - ARPEGE metropole dbl
 observations conventionnelles et satellites
 cumul du DFS sur la période 2015120100 - 2015120118 : 481780



GPS ground	0.63%	SSMIS	3.02%	SYNOP/SYNOR/RADNMF	0.77%
GPS sat	5.71%	GMI	0.00%	SHIP	0.21%
SATOB	7.15%	AIRS	0.69%	PILOT/PRF	1.14%
ATOVS HTRS	0.88%	TAST	20.73%	TFMP	6.59%
ATOVS AMSU-A	12.14%	CRIS	4.46%	AIRCRAFTS	12.89%
ATOVS AMSU-B	6.82%	GEORAD	4.95%	RADAR Vr	0.00%
SAPHIR	2.03%	SCATT	4.96%	RADAR Hvr	0.00%
ATMS	4.46%	BUOY	0.29%	BOGUS	0.00%

Tuning of error variances

- Normalization of \mathbf{R}_i : $s_i^{o,2} \mathbf{R}_i$

$$\begin{aligned} \text{Coef. } s_i^{o,2} \text{ diagnosed with } s_i^{o,2} &= E[J_i^o(\mathbf{x}^a)] / (E[J_i^o(\mathbf{x}^a)])^{\text{opt}} \\ &= E[J_i^o(\mathbf{x}^a)] / (p_i - \text{Tr}(\mathbf{P}_i \mathbf{H} \mathbf{K} \mathbf{P}_i^T)). \end{aligned}$$

- Normalization of \mathbf{B} : $s^{b,2} \mathbf{B}$

$$\begin{aligned} \text{Coef. } s^{b,2} \text{ diagnosed with } s^{b,2} &= E[J^b(\mathbf{x}^a)] / (E[J^b(\mathbf{x}^a)])^{\text{opt}} \\ &= E[J^b(\mathbf{x}^a)] / \text{Tr}(\mathbf{H} \mathbf{K}). \end{aligned}$$

(Desroziers and Ivanov 2001, Chapnik et al, 2004, Desroziers et al 2009)

- Equivalent to Maximum-likelihood estimation (Chapnik et al 2006, Menard 2015)

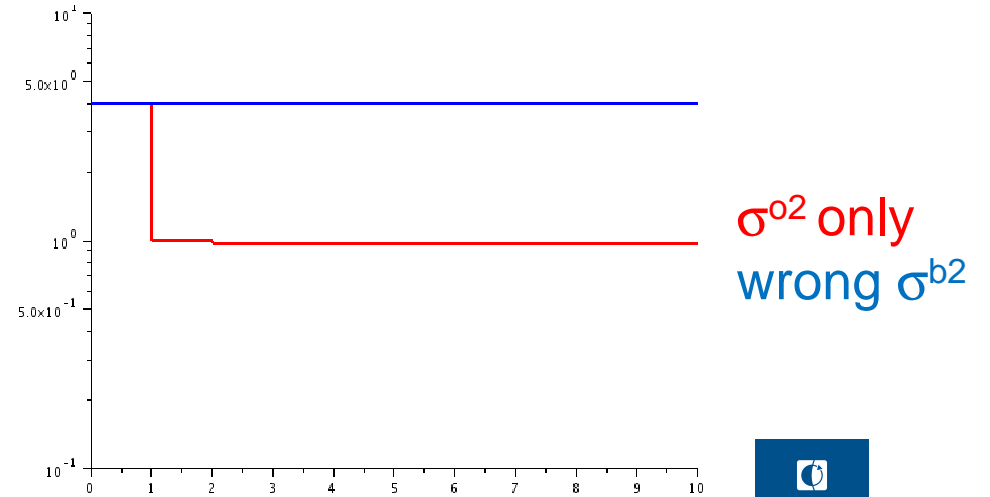
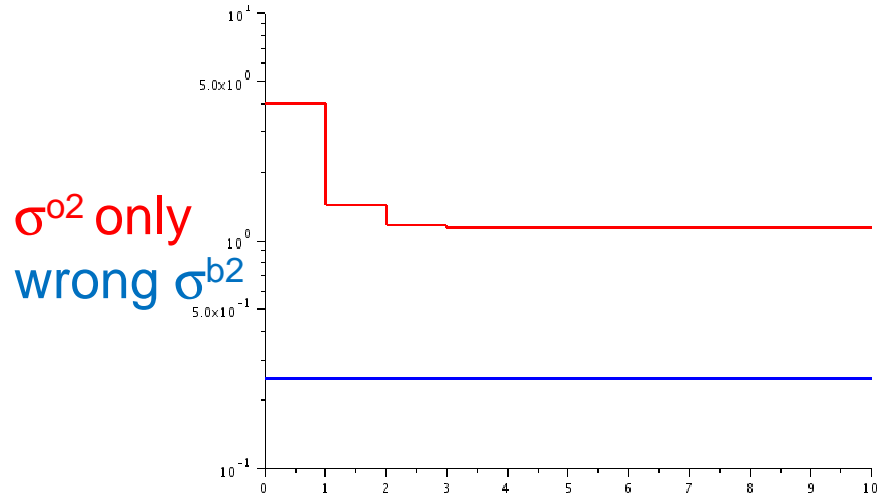
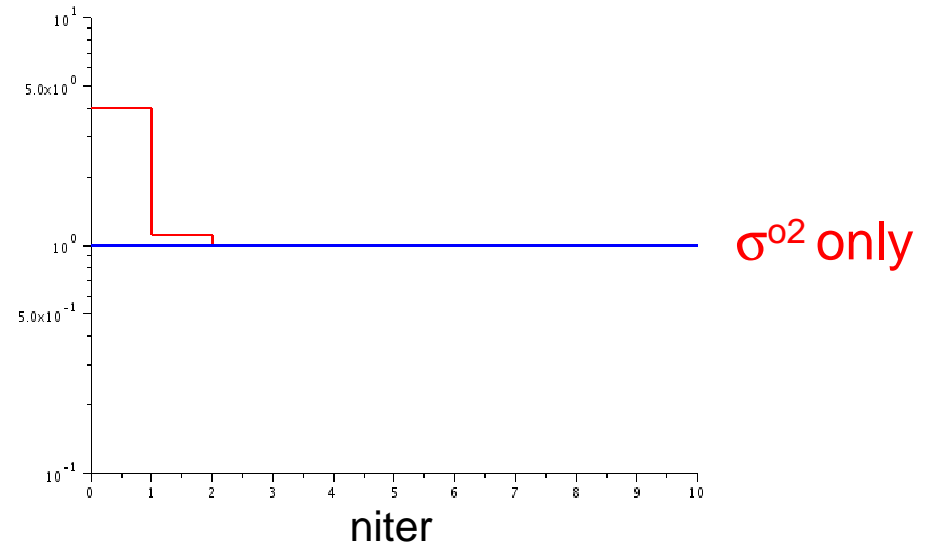
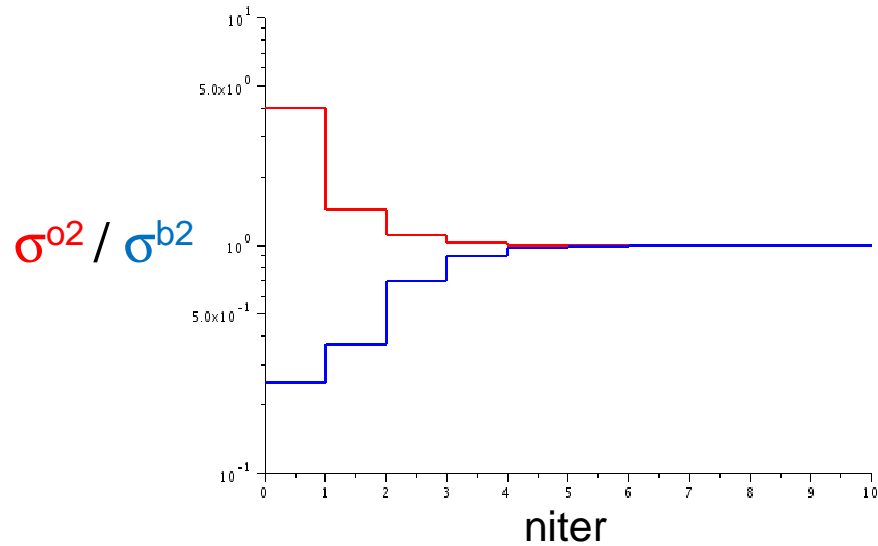
$f(\mathbf{d}|\mathbf{s}) = 1 / ((2\pi)^p \det(\mathbf{D}(\mathbf{s}))^{1/2} \exp(-1/2 \mathbf{d}^T \mathbf{D}(\mathbf{s})^{-1} \mathbf{d}))$,
where $\mathbf{D}(\mathbf{s})$ is the covariance matrix of innovations with parameters \mathbf{s} .

Optimal parameters \mathbf{s} are those that minimize the Log-likelihood
 $L(\mathbf{s}) = -\log (f(\mathbf{d}|\mathbf{s}))$.



J_{\min} diagnostics: mean σ^o / σ^b

$\sigma^{ot} = 1$ ($L^{ot} = 0$ km) / $\sigma^{bt} = 1$ ($L^{bt} = 500$ km)



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Diagnostics in observation space

- Observations / background departures

$$\mathbf{d} = \mathbf{y}^o - H(\mathbf{x}^b).$$

- Observations / analysis departures

$$\begin{aligned}\mathbf{d}^{\text{oa}} &= \mathbf{y}^o - H(\mathbf{x}^a) = \mathbf{y}^o - H(\mathbf{x}^b + \mathbf{K} \mathbf{d}) \\ &\sim (\mathbf{I} - \mathbf{H}\mathbf{K}) \mathbf{d} \\ &\sim \mathbf{R} (\mathbf{H}\mathbf{B}\mathbf{H}^T + \mathbf{R})^{-1} \mathbf{d}.\end{aligned}$$

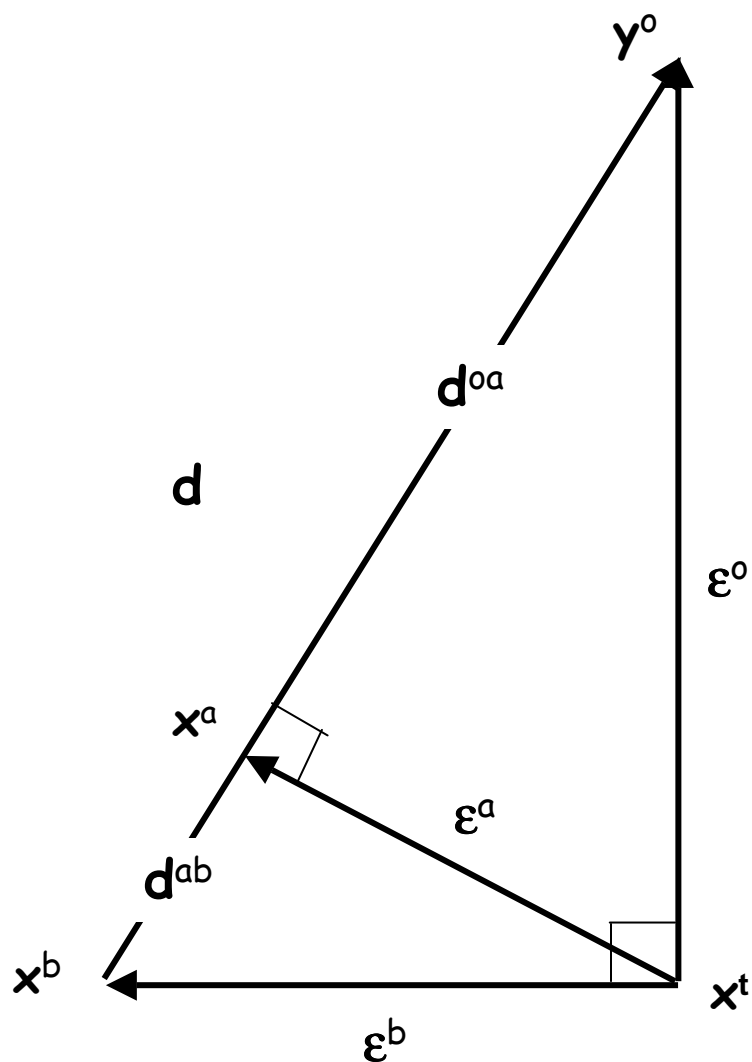
- Background / analysis departures in observation space

$$\begin{aligned}\mathbf{d}^{\text{ba}} &= H(\mathbf{x}^b) - H(\mathbf{x}^a) = H(\mathbf{x}^b) - H(\mathbf{x}^b + \mathbf{K} \mathbf{d}) \\ &\sim \mathbf{H}\mathbf{K} \mathbf{d} \\ &\sim \mathbf{H}\mathbf{B}\mathbf{H}^T (\mathbf{H}\mathbf{B}\mathbf{H}^T + \mathbf{R})^{-1} \mathbf{d}.\end{aligned}$$

- A posteriori diagnostics

$$\begin{aligned}E[\mathbf{d}^{\text{oa}} \mathbf{d}^T] &= \mathbf{R} (\mathbf{H}\mathbf{B}\mathbf{H}^T + \mathbf{R})^{-1} E[\mathbf{d} \mathbf{d}^T] = \mathbf{R} (\mathbf{H}\mathbf{B}\mathbf{H}^T + \mathbf{R})^{-1} (\mathbf{H}\mathbf{B}^t\mathbf{H}^T + \mathbf{R}^t) = \mathbf{R}^t \\ E[\mathbf{d}^{\text{ba}} \mathbf{d}^T] &= \mathbf{H}\mathbf{B}\mathbf{H}^T (\mathbf{H}\mathbf{B}\mathbf{H}^T + \mathbf{R})^{-1} E[\mathbf{d} \mathbf{d}^T] = \mathbf{H}\mathbf{B}\mathbf{H}^T (\mathbf{H}\mathbf{B}\mathbf{H}^T + \mathbf{R})^{-1} (\mathbf{H}\mathbf{B}^t\mathbf{H}^T + \mathbf{R}^t) = \mathbf{H}\mathbf{B}^t\mathbf{H}^T.\end{aligned}$$

Diagnostics in observation space



$$d = y^o - H(x^b)$$

$$d^{oa} = y^o - H(x^a)$$

$$d^{ba} = H(x^b) - H(x^a)$$

$$E[d^{oa} d^T] = R$$

$$E[d^{ba} d^T] = HBH^T$$

$$E[d^{ba} d^{oaT}] = HAH^T$$

$$\langle \epsilon, \epsilon' \rangle = E[\epsilon \epsilon'^T]$$

Diagnosics in observation space: practical implementation

- Difficulty

No multiple statistical realizations of \mathbf{d} , \mathbf{d}^{oa} , \mathbf{d}^{ba} for a given analysis!

- An ergodicity hypothesis is necessary

Average over time, space, subset of observations...

- Practical computation

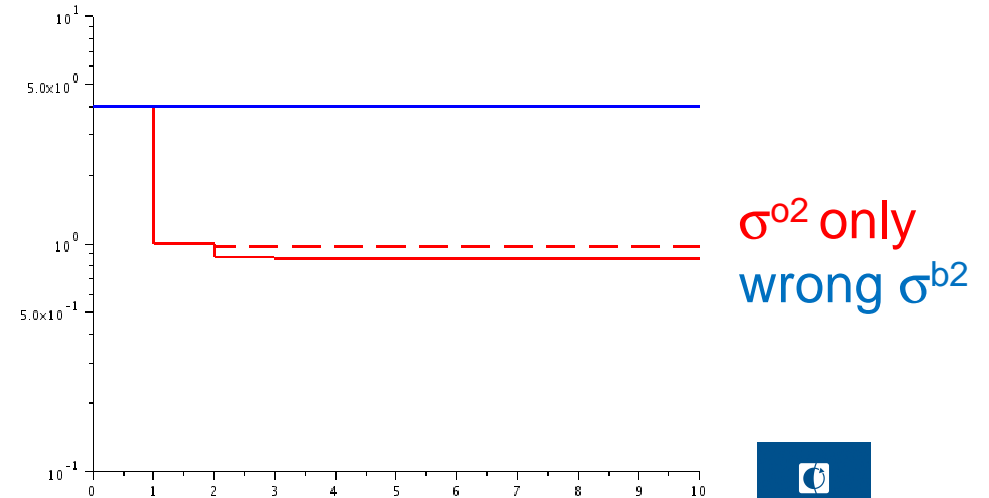
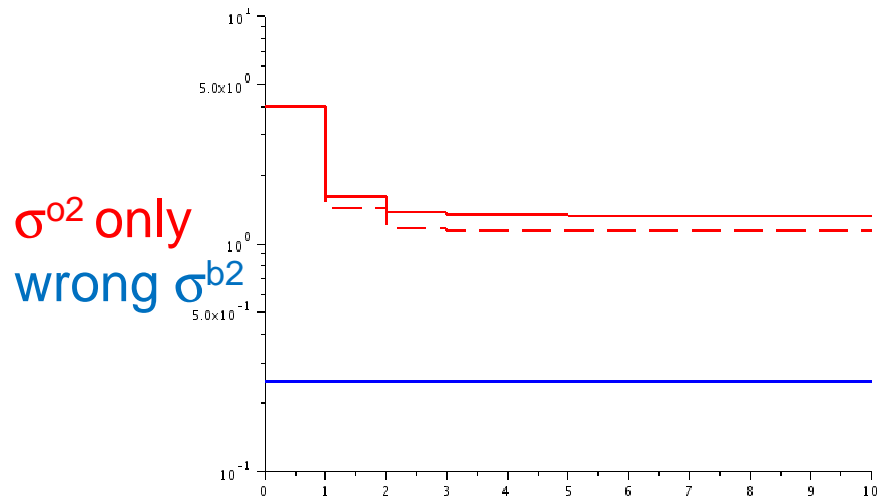
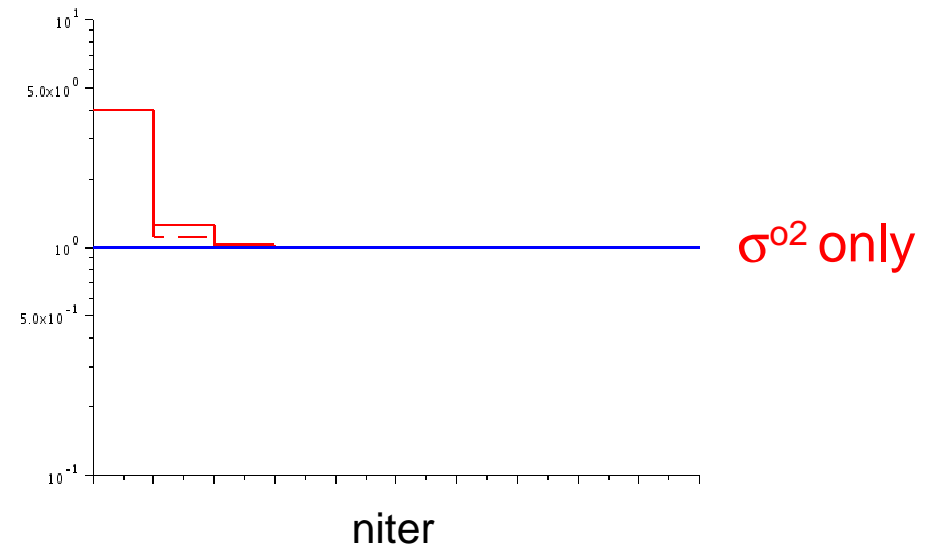
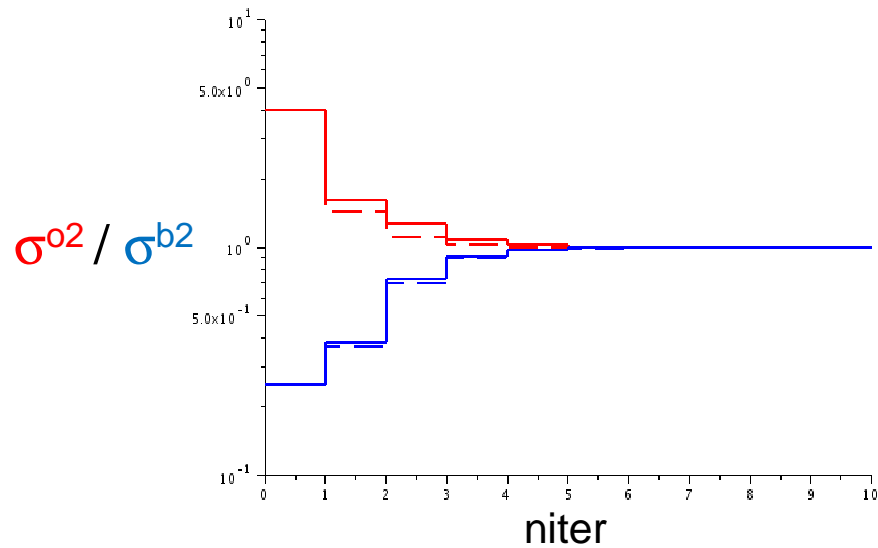
$$(\sigma^{oi})^2 = 1 / p_i \sum_{k=1, p_i} (y_k^{oi} - \mathbf{H}_k^i(\mathbf{x}^a)) (y_k^{oi} - \mathbf{H}_k^i(\mathbf{x}^b)),$$

where p_i is the number of observations in subset i .

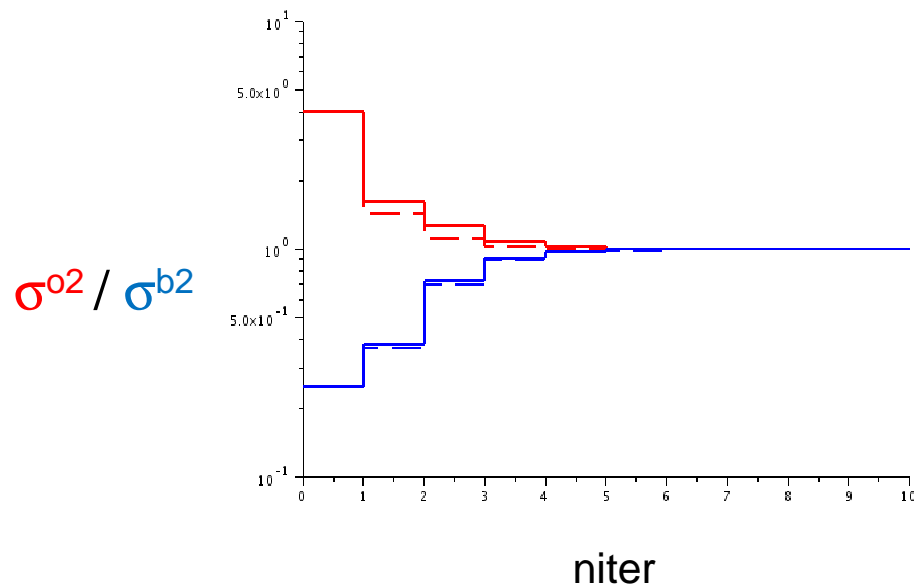
- Easy to implement, especially to try to diagnose observation error covariances!

Diagnosics in observation space: mean σ^o / σ^b

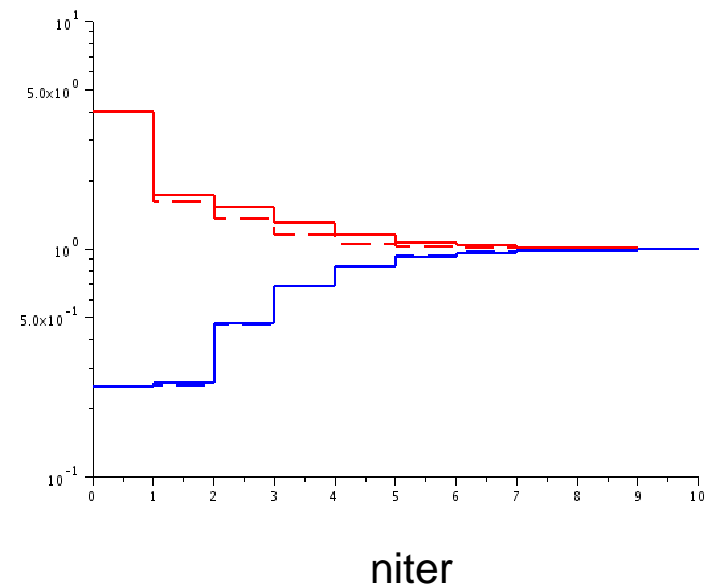
$\sigma^{ot} = 1$ ($L^{ot} = 0$ km) / $\sigma^{bt} = 1$ ($L^{bt} = 500$ km)



Convergence of σ^o , σ^b



$L^b = 500$ km
 $L^o = 0$ km



$L^b = 250$ km
 $L^o = 0$ km

Spectral interpretation

Λ^o specified spectral variances (obs. error)

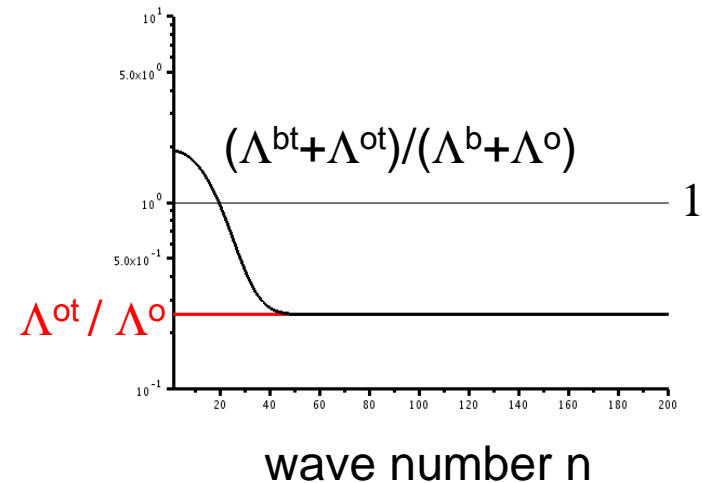
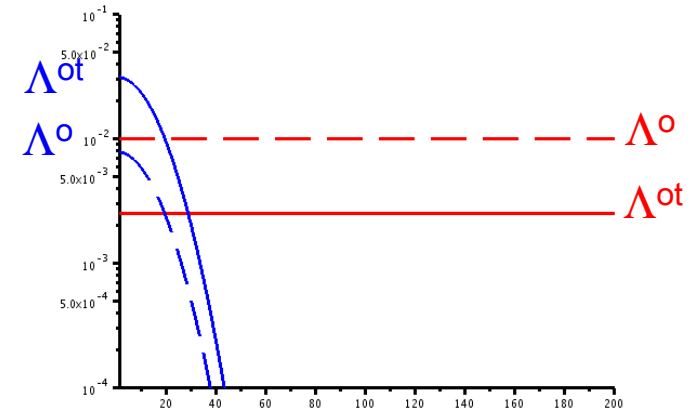
Λ^{ot} true spectral variances

Diagnosed spectral variances:

$$\begin{aligned}\Lambda^{o'} &= \Lambda^o (\Lambda^{bt} + \Lambda^{ot}) / (\Lambda^b + \Lambda^o) \\ &\approx \Lambda^o \Lambda^{ot} / \Lambda^o \\ &\approx \Lambda^{ot}.\end{aligned}$$

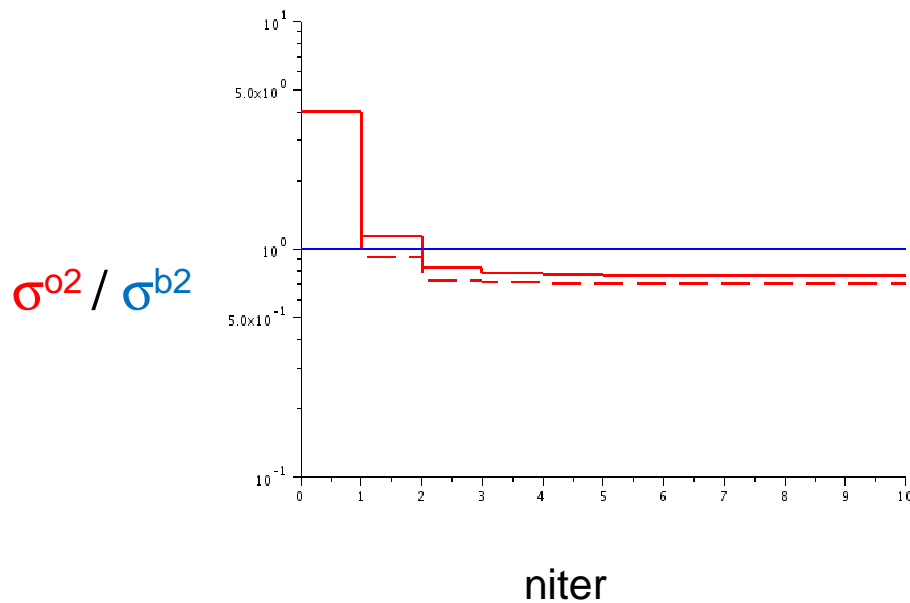
Spatial mean of variance:

$$v^{o'} = \sum_n \Lambda^{o'}(n) \approx v^{ot}.$$

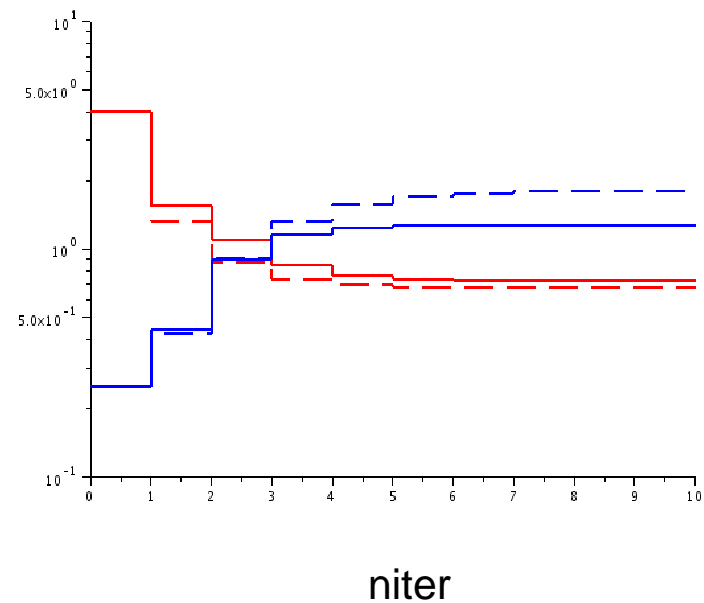


Mis-specification of correlation in R ($L^o = 0$ km)

$\sigma^{ot} = 1$ ($L^{ot} = 100$ km) / $\sigma^{bt} = 1$ ($L^{bt} = 500$ km)



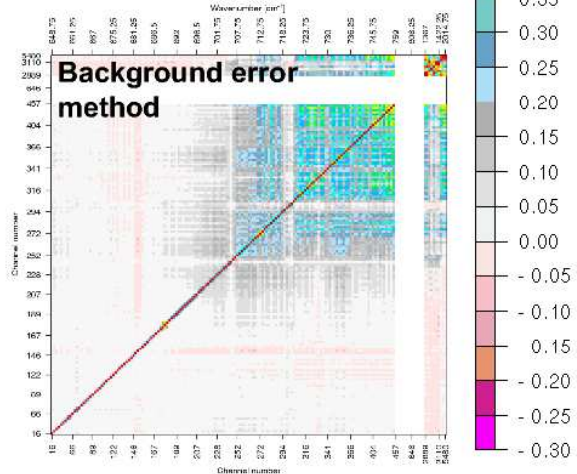
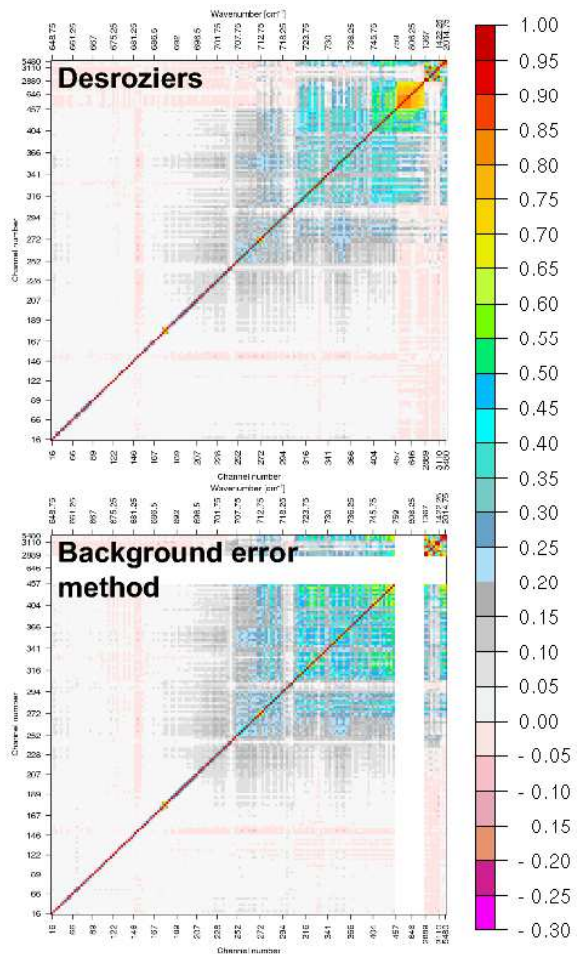
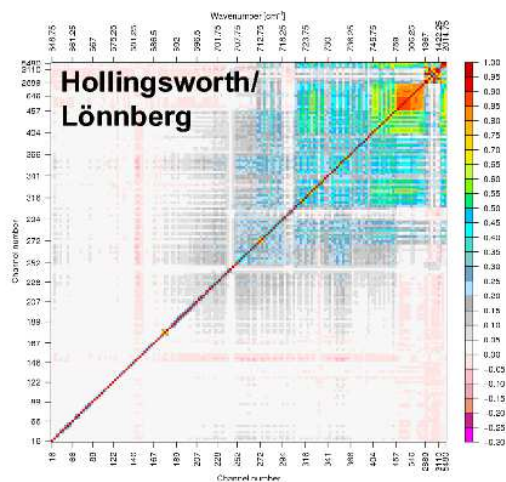
σ^o diagnosis only



σ^o / σ^b diagnosis

IASI inter-channel error correlations

IASI: Inter-channel error correlations



(Bormann et al, ECMWF, 2010)

Diagnostic of model error in 4D-Var [t_{k-1}, t_k]

- Observations / background departures at time t_k

$$\begin{aligned} \mathbf{d}_k &= \mathbf{y}_k^o - H_k(M_k(\mathbf{x}_{k-1}^b)) \\ &\sim \mathbf{y}_k^o - H_k(M_k(\mathbf{x}_{k-1}^t)) - \mathbf{H}_k \mathbf{M}_k \boldsymbol{\varepsilon}_{k-1}^b \\ &\sim \mathbf{y}_k^o - H_k(\mathbf{x}_k^t + \boldsymbol{\varepsilon}_k^m) - \mathbf{H}_k \mathbf{M}_k \boldsymbol{\varepsilon}_{k-1}^b \\ &\sim \boldsymbol{\varepsilon}_k^o - \mathbf{H}_k \boldsymbol{\varepsilon}_k^m - \mathbf{H}_k \mathbf{M}_k \boldsymbol{\varepsilon}_{k-1}^b. \end{aligned}$$

- Observations / analysis departures at t_k with analysis at t_{k-1} but using \mathbf{y}_k^o

$$\begin{aligned} \mathbf{d}_k^{\text{oa}} &= \mathbf{y}_k^o - H_k(M_k(\mathbf{x}_{k-1}^a)) \\ &\sim \boldsymbol{\varepsilon}_k^o - \mathbf{H}_k \boldsymbol{\varepsilon}_k^m - \mathbf{H}_k \mathbf{M}_k \boldsymbol{\varepsilon}_{k-1}^a. \end{aligned}$$

- A posteriori diagnostics

$$\begin{aligned} E[\mathbf{d}_k^{\text{oa}} \mathbf{d}_k^{\text{T}}] &= E[(\boldsymbol{\varepsilon}_k^o - \mathbf{H}_k \boldsymbol{\varepsilon}_k^m) \mathbf{d}_k^{\text{T}}] - E[(\mathbf{H}_k \mathbf{M}_k \boldsymbol{\varepsilon}_{k-1}^a) \mathbf{d}_k^{\text{T}}] \\ &= \mathbf{R}_k + \mathbf{H}_k \mathbf{Q}_k \mathbf{H}_k^{\text{T}} - E[(\mathbf{H}_k \mathbf{M}_k \boldsymbol{\varepsilon}_{k-1}^a) \mathbf{d}_k^{\text{T}}] \\ &= \mathbf{R}_k + \mathbf{H}_k \mathbf{Q}_k \mathbf{H}_k^{\text{T}}, \text{ if optimal } (E[\boldsymbol{\varepsilon}_{k-1}^a \mathbf{d}_k^{\text{T}}] = \mathbf{0}). \end{aligned}$$

(also see Todling 2015)

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Lag-innovation correlation as a diagnostic of optimality

- Innovations used by analysis k-1

$$\mathbf{d}_{k-1} \sim \boldsymbol{\varepsilon}_{k-1}^o - \mathbf{H}_{k-1} \boldsymbol{\varepsilon}_{k-1}^b.$$

- Innovations used by analysis k

$$\begin{aligned} \mathbf{d}_k &= \mathbf{y}_k^o - H_k (M_k (\mathbf{x}_{k-1}^b + \mathbf{K}_{k-1} \mathbf{d}_{k-1})) \\ &\sim \mathbf{y}_k^o - H_k (M_k (\mathbf{x}_{k-1}^t) + \mathbf{M}_k (\boldsymbol{\varepsilon}_{k-1}^b + \mathbf{K}_{k-1} \mathbf{d}_{k-1})) \\ &\sim \mathbf{y}_k^o - H_k (\mathbf{x}_k^t + \boldsymbol{\varepsilon}_k^m + \mathbf{M}_k (\boldsymbol{\varepsilon}_{k-1}^b + \mathbf{K}_{k-1} \mathbf{d}_{k-1})) \\ &\sim \boldsymbol{\varepsilon}_k^o - \mathbf{H}_k \mathbf{e}_k^m - \mathbf{H}_k \mathbf{M}_k (\boldsymbol{\varepsilon}_{k-1}^b + \mathbf{K}_{k-1} \mathbf{d}_{k-1}). \end{aligned}$$

- Covariance of innovations

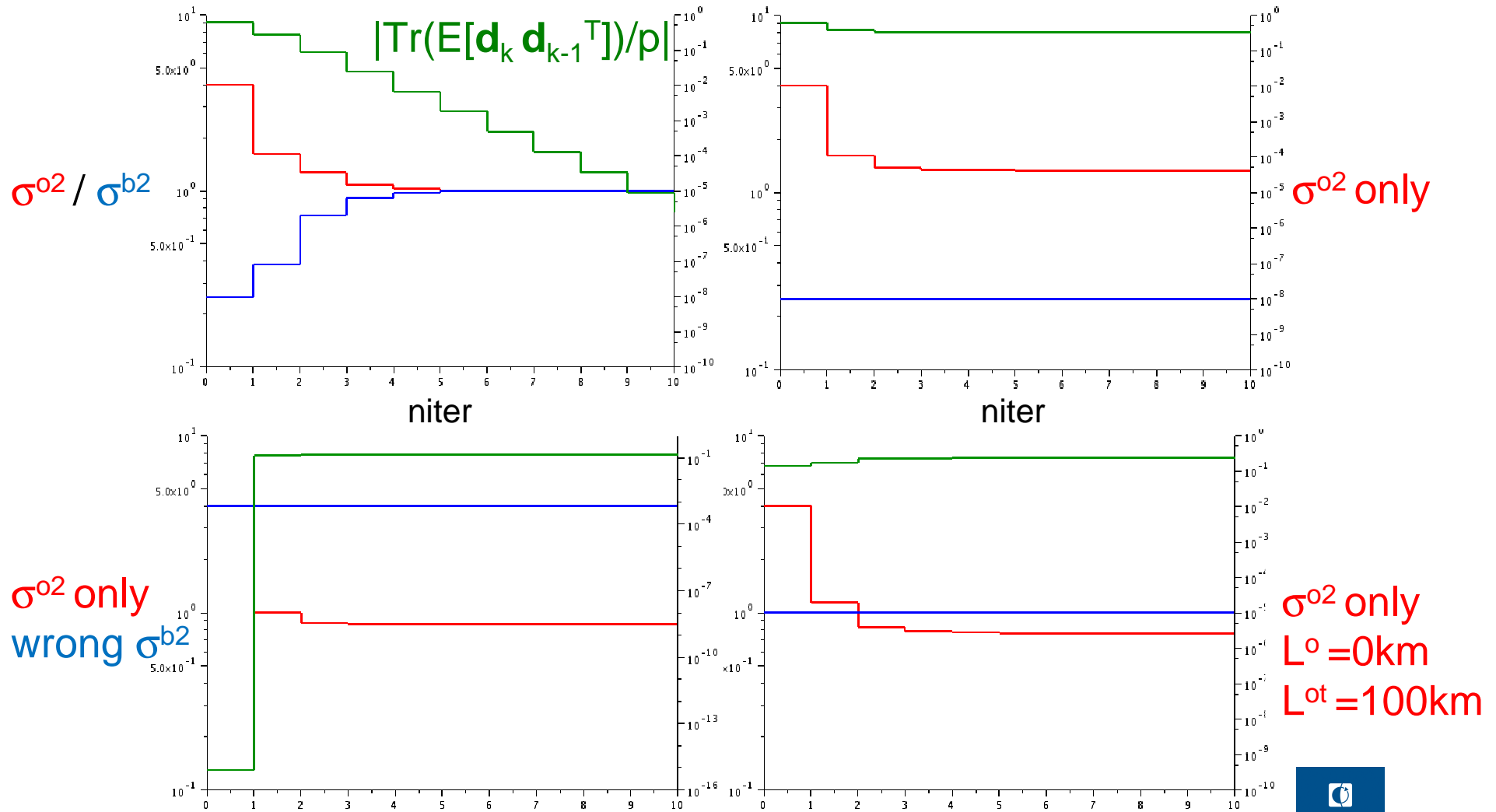
$$\begin{aligned} E[\mathbf{d}_k \mathbf{d}_{k-1}^T] &\sim \mathbf{H}_k \mathbf{M}_k (\mathbf{B}_{k-1}^t \mathbf{H}_{k-1}^T - \mathbf{K}_{k-1} E[\mathbf{d}_{k-1} \mathbf{d}_{k-1}^T]) \\ &\sim \mathbf{H}_k \mathbf{M}_k (\mathbf{B}_{k-1}^t \mathbf{H}_{k-1}^T E[\mathbf{d}_{k-1} \mathbf{d}_{k-1}^T]^{-1} E[\mathbf{d}_{k-1} \mathbf{d}_{k-1}^T] - \mathbf{K}_{k-1}) E[\mathbf{d}_{k-1} \mathbf{d}_{k-1}^T] \\ &\sim \mathbf{H}_k \mathbf{M}_k (\mathbf{K}_{k-1}^t - \mathbf{K}_{k-1}) E[\mathbf{d}_{k-1} \mathbf{d}_{k-1}^T]. \end{aligned}$$

- $E[\mathbf{d}_k \mathbf{d}_{k-1}^T] = 0$

Additional diagnostic of optimality (Ménard 2015)?

Diagnosics in observation space: mean σ^o / σ^b

$\sigma^{ot} = 1$ ($L^{ot} = 0$ km) / $\sigma^{bt} = 1$ ($L^{bt} = 500$ km)



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Conclusion

- J_{\min} diagnostics linked with basic foundations of BLUE estimation
- ✓ Give contributions of observations to the correction of background.
- ✓ Can be used to tune weights of **R** and **B** matrices.

- Observation space diagnostics
- ✓ Rather easy to implement.
- ✓ Well adapted to try to estimate **R**.
- ✓ Give apparently reasonable results in many cases.

- We need something else to validate tunings
- ✓ Impact on forecasts (compared to future independent observations).
- ✓ Compare tuned analysis to independent observations.
- ✓ Use (adapt) lag innovation correlation diagnostic.